

Experiments on Authorship Attribution by Intertextual Distance in English

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Résumé

A quelles conditions peut-on affirmer que des textes sont "proches" ou "éloignés" les uns des autres ? Est-ce que des textes d'un même auteur sont plus similaires entre eux que ceux d'auteurs différents ? La réponse à ces questions est donnée par le calcul de la distance intertextuelle combinée à diverses méthodes de classification. Un test en aveugle et quelques expériences complémentaires montrent que cette méthode offre un outil intéressant pour résoudre les problèmes d'attribution d'auteur par des voies non-traditionnelles.

Abstract

How can it be said that texts are "near" or "distant" from one another? Are different texts by a single author more similar than texts by different authors? To answer these questions, a method is proposed by combination of the calculus of intertextual distance with automatic clustering and tree-classification. A blind test and some additional experiments show that this method offers an interesting tool for non-traditional authorship attribution.

Experiments on Authorship Attribution by Intertextual Distance in English

How can it be said that texts are "near to" or "distant from" one another? Are different texts by a single author more similar than texts by different authors? These are two of the basic questions of authorship (On this question, see: Love, 2002 and for a discussion: Rudman, 1998). Rather than using a set of "frequent words" (Holmes, 1995, Burrows, 2002 & 2003) or "collocations" (Hoover, 2001) or "function words" (Elliott & Valenza, 2001), more or less arbitrarily selected, it would be better to make use of the entire vocabulary of the texts in order to measure proximity. Following this simple idea, a method is proposed to establish authorship by calculating "intertextual distance" (Labbé & Labbé, 2001, 2003).

An index of intertextual distance measures the influence of four factors familiar to literary critics: subject (content, themes), genre, epoch (chronology) and author. A great number of experiments with French texts led to the conclusion that, when looking at texts written during the same period, in the same genre, the predominant factor is nearly always authorship (this tendency seems to be verified also on some English texts: Merriam, 2003b). Thus, this calculation offers an interesting tool for non-traditional authorship attribution.

Tom Merriam applied this method to Marlowe, Shakespeare and Middleton (Merriam, 2002b and 2003a) and some contemporaneous texts (Merriam, 2003b). The results proved encouraging. Thus a question was asked: "Could this technique be extended in English to other authors in the same way that it has been done in French?" Two different experiments gave a first answer to this question.

I. A BLIND-TEST

At the end of January 2004, a blind-test was begun on a corpus of 52 texts — each of 10,000 tokens in length, drawn from English novels — without any indication of the authors' identities or titles of the works (in the following this corpus is named "Oxquarry1"). Texts were supplied by Gerard Ledger who was himself concerned by this question (Ledger, 1995). He chose samples from more or less contemporaneous novels, the authorship of which he thought was difficult to establish by means of their vocabularies. The samples were based on

texts downloaded from the Gutenberg Project. All texts were checked for errors and edited so as to be easily machine readable. However they were not checked against the original works, and only obvious errors were removed or corrected (nevertheless it may be assumed that they are accurate, in that the Gutenberg Project texts are mostly of a high standard). Each excerpt was coded with an alpha-numeric tag in the series 1A-1Z followed by 2A-2Z, giving 52 excerpts. In this article these codes are used throughout, but Appendix 1 discloses authors, titles and chapters of the novels from which these excerpts were drawn. This information was supplied by Dr. Ledger after the following text was sent to him (actual authors and titles are added in order to facilitate the reading of some of the tables and charts).

The question asked was not: "Who wrote what?" but, more modestly: "Is it possible to know which texts are written by the same author and which ones are from different hands?"

To answer this question, a three-step method was used. First, the indices, measuring the distances separating the vocabulary of each pair of texts, were calculated. Secondly, a direct examination of these indices brought out the values which indicate identity of authorship. Thirdly, this separation was complemented by further classifications.

Preliminary remark.

All samples were processed with software written by C. & D. Labbé. The first stage had been designed in parallel with the well-known Oxford Concordance Program (Hockey and Martin, 1988): it lists the words used in the text and calculates their frequencies (the two programs give exactly the same results). Then it calculates the inter-textual distances.

1.1. Calculation of inter-textual distance

Given two texts A and B .

— N_a and N_b sizes of A and B in tokens;

— F_{ia} and F_{ib} the absolute frequencies of type i in texts A and B .

If their lengths (in tokens) are equal — as in this experiment — the distance index can be directly performed by subtracting the two frequencies of each type and by summing the results :

$$D_{(A, B)} = \sum_{i \in (A, B)} |F_{ia} - F_{ib}| \text{ with } N_a = N_b$$

and the relative distance index is equal to:

$$(1) \quad D_{rel(A, B)} = \frac{\sum_{i \in (A, B)} |F_{ia} - F_{ib}|}{2N_a}$$

$D_{(A, B)}$ gives the exact number of different words which separates the two texts. Relative distance index is expressed by a figure between 0 and 1. It could also be defined as "the number of different tokens in an excerpt of 10,000 tokens" and expressed with integers.

This sum is an Euclidian "distance" which gives a quantitative answer to our opening question. Properties of Euclidian distances — identity, symmetry, triangle inequality — allow many useful operations to be performed, especially unbiased classifications (for the properties of textual distance index, see Labbé & Labbé, 2001; Labbé & Labbé, 2003 and below).

It should be noted that this index is given without estimate of uncertainty or error (of course, the spellings of words must have been checked).

The Corpus "Oxquarry1" is 520,000 tokens long. It is sufficient for a first experiment, even if it is not enough to give a full answer to all the questions about authorship attribution. Formula (1) was applied to this corpus. The calculation of all indices leads to a matrix of 52 lines * 52 columns, too large to be printed here. It contains 1,326 different indices (given the identity property, the diagonal matrix is null and given the symmetry property the upper half of the matrix is equal to the lower one).

A direct examination of this matrix gives an initial answer to the authorship question.

1.2. Direct analysis of distance matrix

G. Ledger supplied merely three pieces of information about the corpus Oxquarry1: 1) there were several authors; 2) there was a minimum of two excerpts by each author though not necessarily drawn from the same book; 3) all these texts belonged approximately to the same genre and the same period of time.

Thus if our preliminary assumption — giving authorship a predominant weight as opposed to content — is right, a simple prediction can be made: compared to a random distribution, the distance indices between pairs of excerpts by the same author should fall at the bottom of the distribution, and the highest part of this distribution should contain only pairs of texts by different authors.

The mean of all indices is .4067 and the standard deviation (σ) is equal to .0377. With an uncertainty margin of 5%, we can use two "milestones" (i.e. points on a continuum) to mark out significant values ($\pm 1.96\sigma$ around the mean):

— lower limit (under which the authorship of the two excerpts is probably the same):

$$.4067 - (1.96 * .0377) = .3328 \text{ (rounded to .333).}$$

— upper limit (above which the authorship of the two texts is probably different):

$$.4067 + (1.96 * .0377) = .4806 \text{ (rounded to .481).}$$

a. Same authorship (indices under .333)

70 pairs of texts are under the lower margin (see appendix 2.1). For example, the first four lines of this table indicate that the distance between the nearest couples of texts are:

1D - 2W: .259

1N - 2E: .265

1F - 2E: .267

1K - 2W: .277

Given the transitivity of the inter-textual distance index, it can be said that since:

— texts 1D and 2W are supposed by the same author;

— texts 1K and 2W are also supposed by the same hand;

these two relations imply that the texts 1D and 1K may also be supposed by this same author...

In the same way, it can be said that since pairs of texts 1N-2E and 1F-2E are probably by the same author, similarly texts 1N and 1F may also be supposed by their same author, etc. (in fact, Appendix 1 shows that these excerpts are all by Stevenson).

These relations lead to 12 sets of texts of whose identical authorship is highly probable: this result is arrived at with greater than 95% certainty (Table 1).

Table 1. 12 text sets probably of same authorship

1	1D , 2W, 1K, 2R, 1N, 2E, 1F	Stevenson
2	1S, 2M, 2I, 1C	Morris
3	1E, 1Q, 1B, 2A	Butler
4	2L, 2T, 2J, 2X, 1L, 2D, 1X, 2F	Hardy
5	1H, 2Z, 2S	Hardy
6	2O, 2P, 2V	Forster
7	1Z, 2C, 1Y	Tressel
8	1P, 2U, 1R	Chesterton
9	2H, 1G, 1O	Conrad
10	2N, 2Q, 2K	Conrad
11	1M, 2Y, 1W, 2G	Orczy
12	1J, 2B	Morris

At the end of this first step, five texts remain un-classified: 1A, 1I, 1T, 1U, 1V¹. It must be remembered that the goal was not so much to give a full answer as to avoid a false one.

b. Different authorship (indices above .481)

26 distance indices are above this upper margin (appendix 2.4). These pairs of texts are most probably not by the same author. For example: authorships of 1T and 1Z are different as those of 1T and 2R. Given the fact that:

- 1Z belongs to Set 7, therefore 1T cannot belong to this set 7;
- 2R belongs to Set 1, therefore 1T cannot belong to set 1;

For example, all these mutual exclusions lead to the following conclusions (about the five texts which remained un-classified after examination of the lower limit):

- 1T can only belong to sets 4, 5, 8, 9, 10, 11;
- 1V cannot belong to sets 1 and 12;
- the author of set 10 is different from the authors of sets 1, 2 and 12;

It is worth noting that these negative conclusions are important because, in some cases, it is as valuable to be able to exclude a supposed authorship as to establish an unknown one. Of course, all these deductions are possible, taking into account the information given concerning the corpus (same genre, several authors and, at least, two texts for each author).

¹ 1A : Hardy (Jude); 1I: Orczy (Scarlet P); 1T: Conrad (Almayer); 1U: Orczy (Elusive P); 1V Conrad (Lord Jim).

When a large number of experiments are carried out, we are enabled to establish a standardised distance scale. Our paper (Labbé & Labbé, 2001) presents such a scale for French texts. For example, if, in all experiments to be made on English texts, we never find a distance index under .34 between texts by two different authors, we will be able to assume the same authorship for any two texts when their inter-textual distance is below this threshold (for a 10,000 tokens standardised length).

In order to verify these conclusions and to propose some solutions for unsolved cases, two classification techniques are used.

1.3. Automatic clustering

The Dendrogram below (Fig. 1) gives a pictorial summary of this established analysis.

The algorithm starts by merging the two nearest excerpts (placed at the extreme left of the chart), and then recalculates the distance indices separating this new set from all other texts in order to calculate the next grouping. The nearest text is then placed next to the first two, and the procedure repeated until all texts have been placed in their correct position on the X-axis. Intertextual distance values corresponding to grouping stages are placed on the Y-axis.

The origin of Y-axis is placed at .25 in order to create a kind of "zoom effect" which facilitates reading the diagram, but it must not be forgotten that all the texts are relatively near to another...

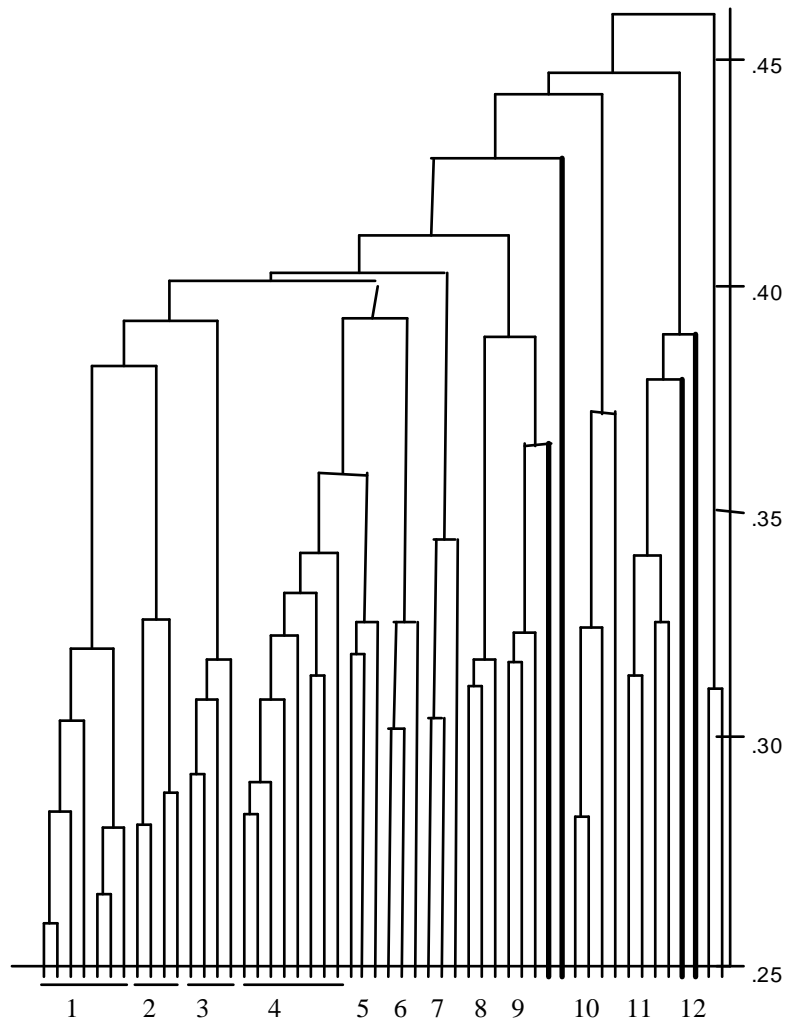
When two texts A , B are joined ($A \cup B$), the properties of intertextual distance allow the calculation, without bias, of a new distance between this group and any text (C) by the arithmetic mean of the original distance indices:

$$(2) D_{(A \cup B, C)} = \frac{D_{(A, C)} + D_{(B, C)}}{2}$$

The calculation is made by merging the two texts A and B and then by calculating the distance index separating this new "text" from C . For example, one can see on the extreme left of the dendrogram, that the horizontal edge between 1D and 2W is placed just under .26 and that this group joins 1K at a level of .275 (mean between indices 1D-1K and 2W-1K)². Thus, horizontal lines which join vertical edges indicate the level of grouping. The higher this level, the more heterogeneous is the group, the more uncertain is the authorship.

² These three excerpts come from Catrionae by Stevenson (Appendix 1).

Fig. 1. Cluster analysis on Corpus Oxquarry1



From left to right :

- | | | | |
|------------|------------|-------------|-------------|
| Set 1 : 1D | Set 4 : 2L | 2C | 2Y |
| 2W | 2T | 1Y | 1W |
| 1K | 2J | Set 8 : 1P | 2G |
| 2R | 2X | 2U | 1U |
| 1N | 1X | 1R | 1I |
| 2E | 1L | Set 9 : 1O | Set 12 : 1J |
| 1F | 2D | 2H | 2B |
| Set 2 : 1S | 2F | 1G | |
| 2M | Set 5 : 1H | 1V | |
| 1C | 2Z | 1A | |
| 2I | 2S | Set 10 : 2N | |
| Set 3 : 1E | Set 6 : 2O | 2Q | |
| 1Q | 2P | 2K | |
| 1B | 2V | 1T | |
| 2A | Set 7 : 1Z | Set 11 : 1M | |

Cluster analysis confirms the conclusions presented above and drawn from the direct examination of the inter-textual distance matrix. It can be seen that all sets have their primary nodes under .35 (calculated by automatic clustering using formula 2). This value is likely to be a threshold for identity of authorship.

In addition, some other conclusions appear to be highly probable (and congruent with the conclusions of direct inspection):

- text 1V may be joined to set 9 (the node is slightly over .35);
- set 4 may be joined to set 5: here again, the node is slightly over .35.

Two other solutions are also suggested and they are congruent with the conclusions of direct inspection³:

— text 1T seems to be near set 10 but this node is far above .35; thus it is impossible to choose between two equivalent solutions: same author, different theme (and/or different book and/or different date) or two different contemporaneous authors and the same theme;

- texts 1U and 1I seem to join set 11 (same uncertainty as with 1T);

In addition, 1A is clearly too far from all the sets and cannot be classified.

A more complete answer is given with the help of a more powerful tool: tree-analysis.

1.4. Tree-classification

The trees below (Figs. 2 and 3) are drawn following Luong's formulae: "valued" trees and "grouping" method (Luong, 1988). These formulae, methods and algorithms are fully explained in Rulhman (2003) and in Labbé & Labbé (2005).

Firstly, this analysis was applied to the original matrix (Fig. 2).

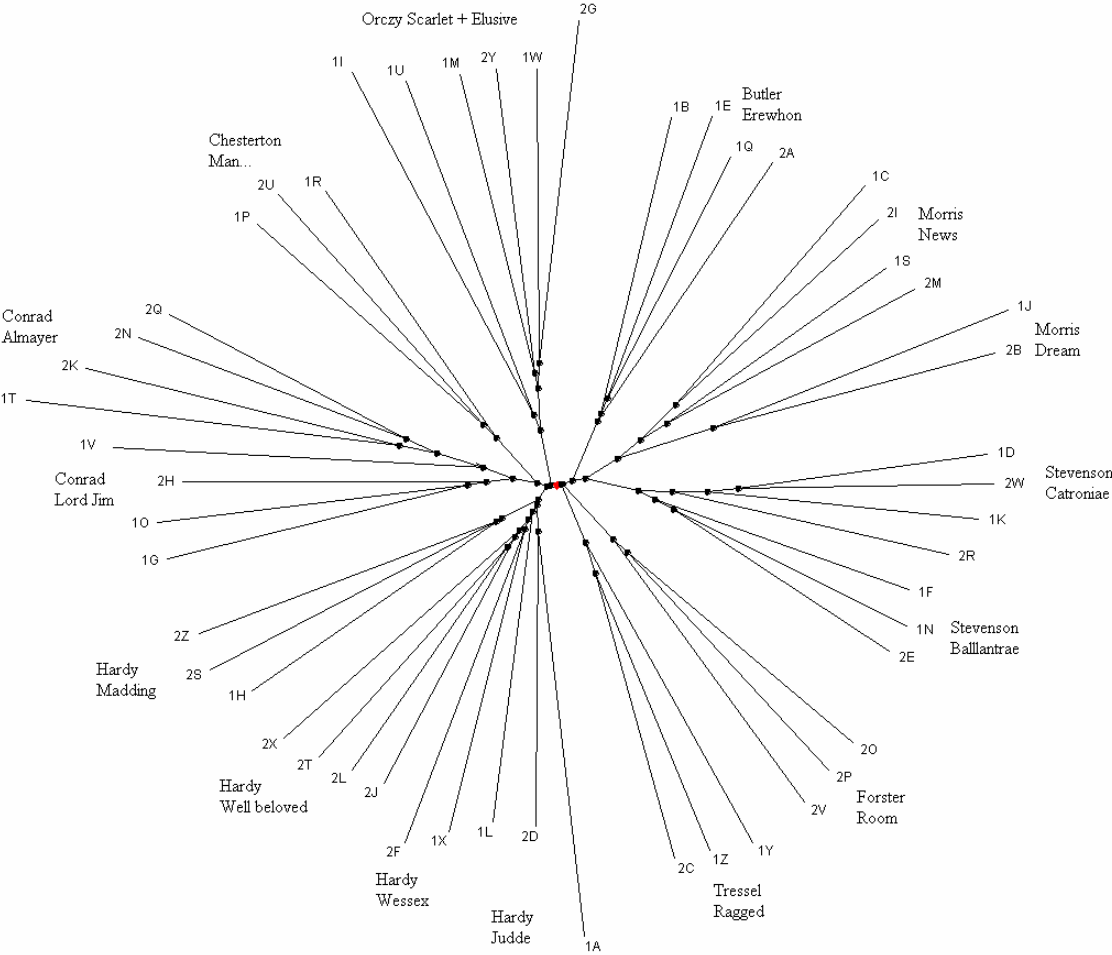
Each "vertice" or terminal "leaf" is tagged with its Corpus code. For example, 1W and 2G (top of the tree), are **adjacent**, as also are 1B and 1E. They form two sets of "**neighbours**" and these two groups are **opposed**. The **edges** ("stems" or "branches") link those four leafs to two **nodes** which are created by the algorithm. Their relative positions are calculated in order to create edges proportional to the original distances. A leaf of the tree is linked to another by a **path** formed by, at least, one edge. The longer the path, the farther apart are the two texts.

This graph is "**valued**". That is to say that the path lengths are positive and in fact proportional to the original values in the corresponding cells of the distance matrix. This

³ In fact, these four solutions are correct: 1V, 1U, sets 9 and 11 are by Orczy; Hardy is the author of Set 4 and 5; Conrad is the one of 1T and set 10 (Appendix 1).

calculation is very complex because this tree must represent, as exactly as possible, the lengths of 1,326 different links.

Fig. 2 Tree-classification on "raw" distance indices



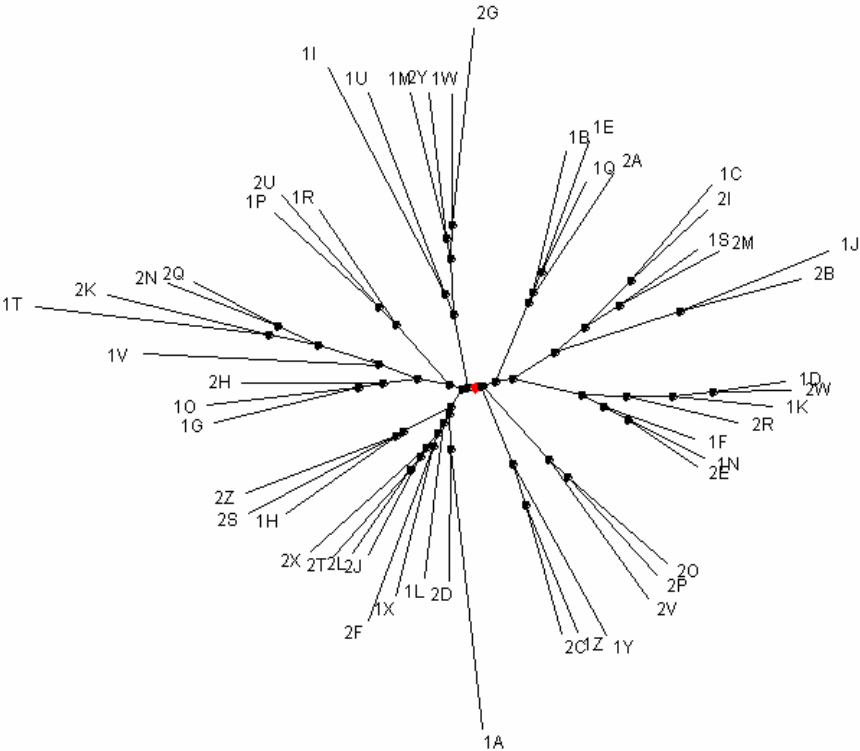
Among various innovations, a measure of **quality** is proposed. The quality of a tree, such as that presented in Fig. 2, can be evaluated by comparing the 1,326 original indices with all corresponding path lengths on the tree. If all these paths are exactly equal to their corresponding distance indices, the quality index will be equal to 1. This index is calculated for each path, for each node and for the whole tree. Above .95, the tree can be considered as an excellent representation of similarities and dissimilarities existing between all the texts. Here the quality index is equal to .975. In other words, 97.5% of the whole information contained in the distance matrix is faithfully represented on this tree. For the 51 nodes, only one index is equal to .934 ; all other indices are over .95. All path quality indices are over .90.

A high degree of confidence can be given to this chart. In all its parts, this tree gives credible information.

Firstly one notices the cross shape (or star shape) around the **tree root**: a small grey lozenge (or red one in electronic files) located exactly at the tree "centre". Terminal edges are very long and nodes, black dots, are often very near the centre of the tree. This shape means that all these texts are roughly "equidistant" except a few pairs - like 1D-2W, 1W-2G (the first nodes are relatively far from the central trunk) - or some texts standing apart (1A, 1T)...

This shape of tree makes it difficult to observe slight differences between texts. Thus the tree can be "reduced" in order to obtain a kind of "zoom effect" on the central trunk of the diagram and to reduce proportionally the length of terminal leaves (in Fig. 3, original distance indices have been "reduced" by 0.2). In other words, this reduction highlights the "contrast" between texts which are actually very near to one another. Of course this "reduction" entails some disadvantages, as it lowers the quality of the tree.

Fig. 3 Tree-classification on distance indices less 0.2



All the conclusions presented above are confirmed and a solution is proposed for the few remaining uncertainties. In addition, tree-classification suggests the reduction of the 12 sets (Table 1) to 9 by using the last junctions before the central trunk (Table 2 below). Texts 1I

and 1U are aggregated to set 1; sets 2 and 12 seem to be by the same author; sets 4 and 5 and text 1A are also assumed to be by one hand, as are sets 9, 10 and texts 1T and 1V (clockwise on Fig. 3).

As G. Ledger confirmed: this classification is correct (in Table 2, author's names have been added in order to facilitate the reading).

Table 2. The nine text sets of probable same authorship. Clockwise on Fig. 2 and 3 (sets already established before tree-classification are placed between brackets)

Final sets	Table 1	Excerpts (between brackets : Table 1)	Authors
1	11	1I, 1U + {1M, 2Y, 1W, 2G }	Orczy
2	3	{1B, 1E, 1Q, 2A}	Butler
3	2 + 12	{1C, 2I, 1S, 2M} + {1J, 2B}	Morris
4	1	{1D , 2W, 1K, 2R, 1F ,1N, 2E}	Stevenson
5	6	{2O, 2P, 2V}	Foster
6	7	{1Z, 2C, 1Y}	Tressel
7	4 + 5	1A + {2D, 1L, 1X, 2L, 2F, 2J, 2T, 2X} + {1H, 2S, 2Z}	Hardy
8	9 + 10	{1G, 1O, 2H} + 1V + {1T, 2K, 2N, 2Q}	Conrad
9	8	1P, 2U, 1R	Chesterton

1.5. Some remarks about this test

What degree of certainty can be assigned to these conclusions?

The sets placed between brackets in table 2 are the only "certain" authorships. In addition, the grouping of the two sets of excerpts by Hardy can be considered as highly probable, because it is clearly indicated by automatic clustering, as well as by tree classification. Other set groupings are "the best fit ones" — taking into account the fact that each author has at least two texts — among all solutions possible, but one cannot assign them a numerical probability. Lengths on the dendrogram, or the tree, give an approximate answer. For example, it can be seen that sets of excerpts by Hardy join together just before reaching the centre of the tree. In Fig. 2, the length of the last edge— joining these sets to the central trunk — is very short: less than 10% of the mean length of the terminal edges of these groups. Of course, the algorithm of classification is accurate but, when differences are so small... it is difficult to choose between two solutions: (1) same author and different themes and/or books;

(2) two contemporaneous authors and a same theme? The same remark applies to pairs {1T - 2K} and {1I - 1U}.

If both cluster and tree classifications have not the same validity as a direct examination of the matrix (with the help of a normalised scale that is yet to be created), their results converge clearly toward the same conclusions. Cluster analysis and tree analysis are therefore reliable tools for literary studies. Further experiments should lead to the calibration of a normalised scale for authorship attribution, like the one used on French texts by Labbé & Labbé (2001 & 2003), which give a standardised degree of assurance.

Is success by chance alone possible?

If there were only pairs of texts, what is the probability of finding all the 26 correct answers by chance? Given the fact that there are 1,326 possible different couples of texts, the probability of success by chance is:

$$P = 26/1,326 * 25/1,325 * (...) * 2/1,301 * 1/1,300$$

This probability is almost infinitesimal. And this figure must be reduced to accommodate possible trios, quartets, etc. Who can easily imagine the number of different combinations that can be made with only 52 objects?⁴

The trial was potentially a difficult one, although the organiser wrote after the trial:

"I ought to point out a concern, which is that critics may object that the texts are relatively easy to differentiate. Some of them should separate into groups quite readily because of certain idiosyncracies, for example the Scottish dialect in the Stevenson work; or the working class dialogue in the Tressel work."

In reply, it is to be noted that, if these idiosyncracies are likely to be prominent details for a human reader, they weigh very little in the whole text surface and in the calculation of the inter-textual distance index. Above all, distances between texts stem from "common" words which are likely to be found in every writer.

Which words? It is an important question to which an answer can be made using Corpus Oxquarry1.

⁴ Here ends the text sent, in February 2004, to G. Ledger in order to present him the results of the "blind-test".

1.6. Contribution of types (vocabulary words) to the inter-textual distances

Where do the distances between texts come from? Or more precisely, two questions may be asked with the help of Corpus Oxquarry1.

First question: Do the differences between texts stem from rare words or from frequent ones?

To answer this question, the software programme calculates, for each pair of texts (A, B), the distance indices between each pair of the same word types which belong to the union of the two texts ($A \cup B$). In the case of the type i ($i \in A, B$), the "absolute contribution to the distance" (between texts A and B) for type i is equal to:

$$D(i_a, i_b) = |F_{ia} - F_{ib}|$$

The surface area of the texts compared in this computation is equal to $F_{ia} + F_{ib}$.

From these can be extracted the "distance index" or "*relative* distance" for the word type i :

$$Drel(i_a, i_b) = \frac{|F_{ia} - F_{ib}|}{F_{ia} + F_{ib}}$$

The "types" may be then ranked by ascending frequency. The contribution of each *relative* frequency group is calculated by summing up the *absolute* contributions (column A in the table 3 below). This allows one to measure the contribution of each group or frequency class to the greater or lesser distancing between the two texts under consideration ($A\%$), as well as the text surface covered by this class (Columns B and $B\%$), and the relative distance (Column C).

Does the value of C differ significantly or not from the average for all the word types? For all of Corpus Oxquarry1, the average or mean distance index is 0.406 ($AvgDrel$) and the standard deviation (σ) is equal to 0.0377. The z -value, or standardised value of the random variable (column D) is equal to:

$$\frac{Drel(i_a, i_b) - Avg(Drel)}{\sigma}$$

Table 3: Contribution by grouped frequency class to the distances (Corpus Oxquarry1)

Frequency Class	A Absolute Distance	A% % of Total Distance	B Theoretical Surface Text	B% % of Total Theoretical Text	C = A/B Distance Index	D Z-value
1 to 4	10,104,730	47.0	12,105,150	22.8	0.835	16.3
5 to 9	2,553,474	11.9	4,480,490	8.4	0.570	6.3
10 to 14	1,233,152	5.7	2,639,652	5.0	0.467	2.4
15 to 19	815,530	3.8	1,986,830	3.7	0.410	0.2
20 to 29	1,098,062	5.1	2,940,266	5.5	0.373	-1.2
30 to 49	1,096,182	5.1	3,426,642	6.5	0.320	-3.2
50 to 99	1,340,886	6.2	5,066,190	9.6	0.265	-5.3
100 to 199	1,202,394	5.6	5,335,286	10.1	0.225	-6.8
200 to 499	1,242,266	5.8	7,132,326	13.4	0.174	-8.8
500 to 999	588,360	2.7	5,066,960	9.6	0.116	-11.0
>999	220,334	1.0	2,853,986	5.4	0.077	-12.5
Sum	21,495,370	100.0	53,033,778	100.0	0.406	

Within the z -value interval (-1.96 et + 1.96), we cannot say that an observed value significantly differs from the overall mean of .406 for all the words in Corpus Oxquarry1. Above and below these two limits, however, there is less than 1 chance in 20 of erring in stating that the observed value differs materially from the overall mean value.

For example, the least frequent "types", those of between 1 and 4 occurrences, contribute 10,105,000 words, that is to say, 47% of the total distance. At the same time, they cover only 22.8% of the total "surface". Therefore the relative distance or index is equal to 0.835 (10,104,730/12,105,150). The z -value of +16.3 indicates that there is virtually no chance of this being the result of random variation; *the "rare words" are those which create the greatest distances between the texts*. By contrast, the various frequent words (>499) contribute only 3.7% to this distance, while they occupy some 15% of the surface.

In referring to the article which appeared in *Corpus* (Labbé & Labbé, 2003, Table 4, p. 113), one will recognize that, even though the profile is the same as in the table above, the z -values are less in the case of a similar corpus of French works. In effect, lemmatization in French reduces the weighting of low frequency words and proportionally increases that of the overall average and high frequency words.

These contrasts suggest some important comments: For example, most linguistic analyses – notably PCA and factor analysis – are concerned only with the most frequent "types", whereas it is manifestly the rare words which are the principal differentiators of authors.

Second question : among the frequent word types, which ones do and which ones do not differentiate between the texts?

Here the concern is only with the most frequent "types". One already knows that on average their contribution is weak. NB: The list of these words is based on the *entire* Corpus Oxquarry1 (the first column of the two tables 4 and 5 below gives the frequency of these "types" for the entire corpus). The reasoning is the same as above. The "types" are ranked according to their ascending contribution (i.e., increasing distance index). Tables 4 and 5 show the top and bottom of the list.

Table 4: The most frequent "types" (or function words) which contribute least to the distance between texts.

Type	Frequency	Distance	Text Surface	Distance index	Z-value
to	13 147	99 366	1 347 726	0.074	-12.60
the	29 654	231 374	3 036 642	0.076	-12.51
in	8 351	82 456	856 188	0.096	-11.75
of	15 975	167 632	1 635 468	0.103	-11.51
that	6 805	74 808	697 272	0.107	-11.33
and	16 441	181 310	1 684 326	0.108	-11.32
for	4 248	47 272	435 132	0.109	-11.28
at	3199	38 500	327 828	0.117	-10.94
with	4 296	56 604	439 824	0.129	-10.52
a	12 144	162 178	1 245 114	0.130	-10.46
out	1 287	17 254	132 090	0.131	-10.44
was	7 352	100 688	751 944	0.134	-10.32
on	3 126	43 358	320 586	0.135	-10.27
there	1 841	27 330	189 414	0.144	-9.92
as	4 448	67 982	456 246	0.149	-9.74
not	3 486	55 558	356 286	0.156	-9.48
from	1 863	29 884	190 332	0.157	-9.44
this	2 263	36 442	231 846	0.157	-9.43
if	1 630	26 866	167 382	0.161	-9.31
what	1 597	26 538	163 914	0.162	-9.25
but	3 387	56 818	346 902	0.164	-9.18

Table 5: The frequent "types" which contribute most to the distance between texts

Type	Frequency	Distance	Text Surface	Distance index	Z-value
your	977	40 654	100 266	0.4055	0.01
house	474	20 320	48 552	0.4185	0.50
things	350	15 092	35 904	0.4203	0.57
light	300	13 476	30 804	0.4375	1.22
yes	398	18 172	41 412	0.4388	1.27
moment	320	14 528	32 640	0.4451	1.51
felt	303	14 350	30 906	0.4643	2.24
my	2 529	122 838	260 406	0.4717	2.52
we	1 462	71 168	149 532	0.4759	2.69
rather	385	19 064	39 168	0.4867	3.10
us	511	25 564	52 224	0.4895	3.20
men	592	29 966	60 690	0.4938	3.36
woman	324	16 430	33 150	0.4956	3.43
our	519	27 104	53 040	0.5110	4.02
don't	497	26 444	51 204	0.5164	4.22
she	3 731	200 024	380 868	0.5252	4.56
her	3 977	214 302	406 062	0.5278	4.66
people	456	25 450	46 614	0.5460	5.35
shall	617	39 172	63 444	0.6174	8.06
work	318	20 596	32 436	0.6350	8.73
sir	312	22 484	32 436	0.6932	10.94
Mr	698	51 892	71 808	0.7227	12.06
Mrs	334	26 118	34 170	0.7644	13.65
father	368	28 912	37 740	0.7661	13.71

As in French, prepositions, articles, conjunctions, and auxiliary verbs produce the smallest distances between texts. In other words, they are common to all authors, or, similarly, the structure of the language imposes rigid restraints on those who use it; therefore authors can hardly be differentiated by means of these words. Individual differences are therefore, on average, very weak.

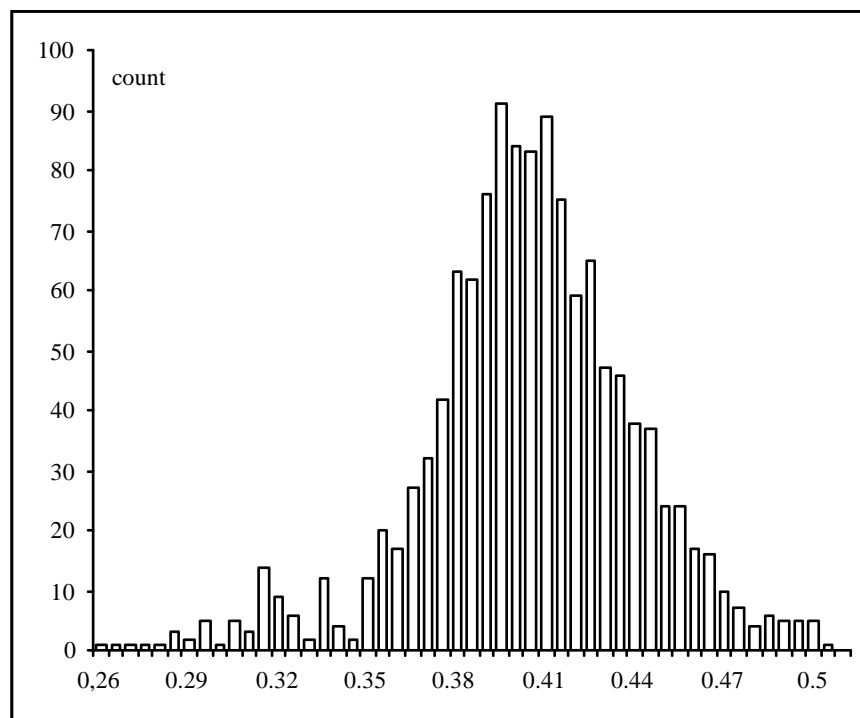
By contrast, personal pronouns, verbs, and the most ordinary nouns (substantive, adjectives) are more variable in their usage among different texts and authors. That is the case in French as well.

A main conclusion can be drawn from the analysis of distance contributions: for authorship attribution, a great number of experiments like this may lead to a set of more or less frequent words able to discriminate authors in general. Texts of different length will then become comparable. One can understandably be sceptical about this outcome, but one can imagine how to validate the idea, applying the procedure to a great number of texts which are of established authorship. Then one could apply these lists to unknown or doubtful texts. It would be a good supplement to a standardised scale for authorship attribution.

1.7. A provisional scale for authorship attribution

The reasoning above postulates that the indices of inter-textual distance calculated between texts by a same author must be significantly lower than those between texts by different authors. The results of the blind-test appear to validate this postulate but nevertheless a more detailed examination is of interest. Given the choices made by G. Ledger, there are 153 different pairs of excerpts by the same authors and 1173 pairs which group excerpts by different authors. Let the entire 1,326 indices be ranked by ascending values. Appendix 2 reproduces the beginning and the ending tails of this distribution. Fig. 4 presents their repartition in equal classes (absolute densities).

Fig. 4 Distribution of inter-textual distance indices ranked by ascending values in equal classes from Oxquarry1.



The shape of this graph suggests clearly a mix of different populations :

— in the middle of the graph, a large population, with a "mode" around 0.395-0.410 seems to be more or less "Gaussian";

— on the right tail, a little group of text pairs separated by distance indices over 0.48. Appendix 2.4 highlights the singularity of Conrad and suggests that, for some these texts, a supplementary factor (genre?) adds its effects to differences in authorship and content.

— on the left tail, two small populations of which one cannot say if their distribution is Gaussian or not because they are of insufficient size. The first population from the left has a mode at 0.315-0.320; the second population mode seems to take place at 0.325-0.330. Appendix 2.1 indicates clearly the difference between these two populations. The first one is quite exclusively composed of text pairs drawn from the same books; the second one contains a majority of excerpts from two different books by the same author.

Of course these populations are not strictly separated, but some thresholds appear clearly:

— all indices less than .349 are within excerpts by the same author. 56% of all distances between texts by the same author are below this threshold (Appendix 2.1).

— between .35 and .36 one can find a great majority of pairs of texts by the same author (Appendix 2.2, marked in bold) and very few of them can be found over .36 and under .395 (Appendix 2.3).

— all indices above .395 are between excerpts by different authors : 78% of all pairs of texts by two different authors are above this threshold.

It is also noticeable that the highest indices (Appendix 2.3) observed for the same author, concern for most part only different books, except Orczy's Scarlet Pimpernel (.364) and Elusive P. (.376), Conrad's Lord Jim (.367 and 373) and Almayer's Folly (.384). For these four books, various explanations are possible, for example: they are more "heterogeneous" than the others; G. Ledger chose excerpts very different from each other.

A final remark: in Corpus Oxquarry1, the texts are of equal lengths. Are the results equally good when the method is applied to texts the lengths of which are different? In other words, what is the influence of text length on inter-textual distance? And, if there is an actual influence, how could one neutralise this drawback?

II. INFLUENCE OF TEXT LENGTHS ON DISTANCE INDEX

Firstly, a modification of formula (1) is presented to adapt it in the case of texts unequal in length. To measure the potential bias due to the differences of length, several trials were carried out. These experiments followed the same principle: the length was the only factor allowed to vary while the other factors were neutralised (as strictly as possible).

2.1. Calculation of Inter-textual distance index on texts of different sizes

The extended reasoning is presented in Labbé & Labbé (2001 and 2003). It is proposed to "reduce" the larger text (B) to the size of the shorter one (A) as explained below. Let N_a and N_b be the size of the two texts (in tokens), and a type i of which absolute frequencies are f_{ia} and f_{ib} in A and B . The mathematical expectancy of this type in A , considering its frequency in B (f_{ib}) is:

$$E_{ia(u)} = F_{ib} * u \text{ with } u = \frac{N_a}{N_b}$$

and the "absolute" distance is: $F_{ia} - E_{ia(u)}$

Thus, the new distance index between A and B is obtained in substituting $E_{ia(u)}$ for F_{ib} in the formula (1):

$$(3) \quad D_{rel(A, B)} = \frac{\sum_{i \in (A, B)} |F_{ia} - E_{ia(u)}|}{2N_a}$$

Two main experiments were carried out in order to test the accuracy of the formula (3) and to measure its possible sensitivity to the differences of text lengths. First, the Corpus Oxquarry1 was used, then a larger one, Oxquarry2, was collected with the help of G. Ledger.

2.2. Experiment on short texts (Oxquarry1)

Given three texts A , B , C , when two of them (A , B) are joined ($A \cup B$), is the new distance index between this group and any other text (C), equal to the arithmetic mean of the original distances (see formula (2) above)? In what follows, the indices calculated in applying formula (2) are designated "*theoretical*", and those using formula (3) on the actual "new text" created by joining A to B are designated "*observed*". If the results of the two calculations are equal, it means that the calculation is insensitive to differences in text lengths.

In the blind-text presented above,

— the two nearest texts are excerpts from Stevenson's Catrionae (1D and 2W). These two excerpts were joined and this "new text" was designated "Stevenson2" (since the length of this "text" is 20,000 tokens, the double of the other excerpts which are unchanged);

— the most distant excerpts written by a single author are 1I (Orczy, Elusive) and 1U (Orczy, Scarlet). They were also joined ("Orczy2").

The theoretical and observed indices are calculated between these two "new" texts (20,000 tokens long) and the 48 remaining excerpts (10,000 tokens long). Table 6 below gives the results of the comparison between the two results.

This comparison reveals:

— a systematic bias: observed indices are always less than the theoretical ones. Almost all the discrepancies vary between 2% and 4%, but they are above 5% in the case of two occurrences for Stevenson and in four for Orczy (bold characters in Table 6). Given the fact that for French texts of different lengths, the results fall within an uncertainty margin of 5%, these six indices must be considered as "anomalous".

It is worth noting that, for Orczy, half of the discrepancies come from distances within her own work (between Elusive P. and Scarlet P.). More generally, the highest values in the last two columns are observed between texts by the same author or between the nearest texts by different authors.

Table 6 Influence of text size on distance indices (Corpus Oxquarry1, scale difference 1:2)

	Theoretical index		Observed index		Difference %	
	Stevenson2	Orczy2	Stevenson2	Orczy2	Stevenson2	Orczy2
1A (Hardy Jude)	0.4384	0.4228	0.4267	0.4057	-2.7	-4.0
1B (Butler Erewhon revisit.)	0.4164	0.4394	0.4081	0.4265	-2.0	-2.9
1C (Morris News)	0.4122	0.4355	0.4034	0.4219	-2.1	-3.1
1E (Butler Erewhon revisit.)	0.3962	0.4424	0.3836	0.4299	-3.2	-2.8
1F (Stevenson Ballantrae)	0.3315	0.4058	0.3158	0.3857	-4.7	-5.0
1G (Conrad Lord Jim)	0.3879	0.4343	0.3776	0.4171	-2.7	-4.0
1H (Hardy Madding)	0.4138	0.4145	0.4037	0.3969	-2.4	-4.3
1J (Morris Dream of JB)	0.4256	0.4560	0.4143	0.4418	-2.7	-3.1
1L (Hardy Jude)	0.2842	0.4378	0.2644	0.4231	-7.0	-3.4
1M (Orczy Scarlet P)	0.4052	0.4197	0.3957	0.4027	-2.4	-4.1
1N (Stevenson Ballantrae)	0.4353	0.4034	0.4268	0.3841	-2.0	-4.8
1O (Conrad Lord Jim)	0.3256	0.4237	0.3106	0.4056	-4.6	-4.3
1P (Chesteron Man who was)	0.4290	0.4304	0.4194	0.4153	-2.2	-3.5
1Q (Butler Erewhon revisit.)	0.4373	0.4207	0.4288	0.4036	-1.9	-4.1
1R (Chesteron Man who was)	0.3915	0.4327	0.3810	0.4173	-2.7	-3.6
1S (Morris News)	0.4059	0.4155	0.3963	0.3980	-2.4	-4.2
1T (Conrad Almayer)	0.3572	0.4279	0.3459	0.4138	-3.1	-3.3
1V (Conrad Lord Jim)	0.5118	0.4612	0.5029	0.4456	-1.7	-3.4
1W (Orczy Elusive P)	0.4764	0.4365	0.4679	0.4187	-1.8	-4.1
1X (Hardy Wessex Tales)	0.4346	0.3925	0.4242	0.3705	-2.4	-5.6
1Y (Tressel Ragged TP)	0.4157	0.4173	0.4037	0.3991	-2.9	-4.4
1Z (Tressel Ragged TP)	0.4413	0.4249	0.4319	0.4050	-2.1	-4.7
2A (Butler Erewhon revisit.)	0.4325	0.4295	0.4215	0.4148	-2.5	-3.4
2B (Morris Dream of JB)	0.4177	0.4299	0.4086	0.4127	-2.2	-4.0
2C (Tressel Ragged TP)	0.4330	0.4290	0.4246	0.4151	-1.9	-3.2
2D (Hardy Jude)	0.4184	0.4263	0.4100	0.4115	-2.0	-3.5
2E (Stevenson Ballantrae)	0.4148	0.4064	0.4038	0.3871	-2.7	-4.7
2F (Hardy Wessex Tales)	0.3244	0.4148	0.3112	0.3958	-4.1	-4.6
2G (Orczy Elusive P)	0.4284	0.4306	0.4169	0.4138	-2.7	-3.9
2H (Conrad Lord Jim)	0.4599	0.3977	0.4501	0.3747	-2.1	-5.8
2I (Morris News)	0.4033	0.4181	0.3932	0.4006	-2.5	-4.2
2J (Hardy Well beloved)	0.3902	0.4334	0.3793	0.4192	-2.8	-3.3
2K (Conrad Almayer)	0.4057	0.4138	0.3948	0.3967	-2.7	-4.1
2L (Hardy Well beloved)	0.4861	0.4379	0.4773	0.4227	-1.8	-3.5
2M (Morris News)	0.3983	0.4237	0.3877	0.4065	-2.6	-4.1
2N (Conrad Almayer)	0.3810	0.4260	0.3686	0.4125	-3.3	-3.2
2O (Forster Room with view)	0.4546	0.4224	0.4474	0.4055	-1.6	-4.0
2P (Forster Room with view)	0.3977	0.4316	0.3878	0.4139	-2.5	-4.1
2Q (Conrad Almayer)	0.4245	0.4235	0.4135	0.4055	-2.6	-4.2
2R (Stevenson Catrionae)	0.4339	0.4207	0.4265	0.4037	-1.7	-4.0
2S (Hardy Madding)	0.2966	0.4154	0.2788	0.3949	-6.0	-4.9
2T (Hardy Well beloved)	0.4076	0.4163	0.3962	0.3990	-2.8	-4.2
2U (Chesteron Man who was)	0.4240	0.4182	0.4136	0.3989	-2.4	-4.6
2V (Forster Room with view)	0.4293	0.4249	0.4196	0.4071	-2.2	-4.2
2W (Stevenson Catrionae)	0.4098	0.4395	0.3992	0.4211	-2.6	-4.2
2X (Hardy Well beloved)	0.4336	0.4264	0.4240	0.4096	-2.2	-3.9
2Y (Orczy Scarlet P)	0.4386	0.3668	0.4286	0.3400	-2.3	-7.3
2Z (Hardy Madding)	0.4226	0.4154	0.4111	0.3969	-2.7	-4.5
Mean					-2.5	-4.2

The same experiment was repeated:

— step 3: the closest excerpt was added to Stevenson2 (i.e. 1K which is also an excerpt from Catrinae) in order to constitute "Stevenson3" and to Orczy2 (1M, Scarlett P.) in order to constitute "Orczy3";

— step 4: 2R is added to Stevenson3 (Stevenson4) and 1W to Orczy3 (Orczy4) ;

— steps 5 & 6. Stevenson5 and 6 are formed by merging with Stevenson4: 1F and then 1N; Orczy5 and 6 are constituted by merging with Orczy4: 2G and then 2Y.

It is not possible to proceed further in this way because there are only six excerpts by Orczy. The distance indices between these groups and the remaining 10,000 token length excerpts were calculated. This experiment highlights index behaviour when calculated with texts the sizes of which diverge further and further. Tables 7 and 8 below show the results of this new experiment.

Table 7 Stevenson. Comparison with all remaining 10,000 token excerpts
Means of the whole indices

	A Means of Observed Indices	B Means of Theoretical Indices	(A-B)/B%
Stevenson 20,000	.402	.412	-2.5
Stevenson 30,000	.397	.413	-3.9
Stevenson 40,000	.379	.412	-8.0
Stevenson 50,000	.374	.406	-7.9
Stevenson 60,000	.372	.405	-8.2

Table 8 Orczy. Comparison with all remaining 10,000 token excerpts
Means of the whole indices

	A Means of Observed Indices	B Means of Theoretical Indices	(A-B)/B%
Orczy 20,000	.406	.424	- 4.2
Orczy 30,000	.390	.421	-7.3
Orczy 40,000	.387	.420	-7.9
Orczy 50,000	.392	.425	-7.9
Orczy 60,000	.391	.426	-8.2

The systematic bias noted above appears clearly. It is also clear that the relation between text lengths and distance indices is non-linear. As the differences of size between texts increase, the growth rate of the divergences between theoretical and observed values decreases (last column). The same phenomenon was noted with the French corpora but to a

lesser degree (Labbé & Labbé, 2003). Again, these differences stem from a lesser proportion of low frequency word types in French corpora (due to lemmatisation).

This bias entails an obvious consequence: when comparing texts the sizes of which greatly differ, the indices calculated in applying Formula 3 to the longest ones will tend to be lower than those calculated with the shortest ones.

However, two objections may be made to this experiment: the excerpts by Stevenson are much nearer one another than those by Orczy, and it can be added that the larger the "artificial" texts, the lesser the others : they remain 48 when length factor is equal to 2; 40 when it reaches 6 (it is the reason why values in column B vary).

In order to organise a more systematic experiment, G. Ledger was asked to set up a new corpus.

2.3 First experiment on long texts (Corpus Oxquarry2)

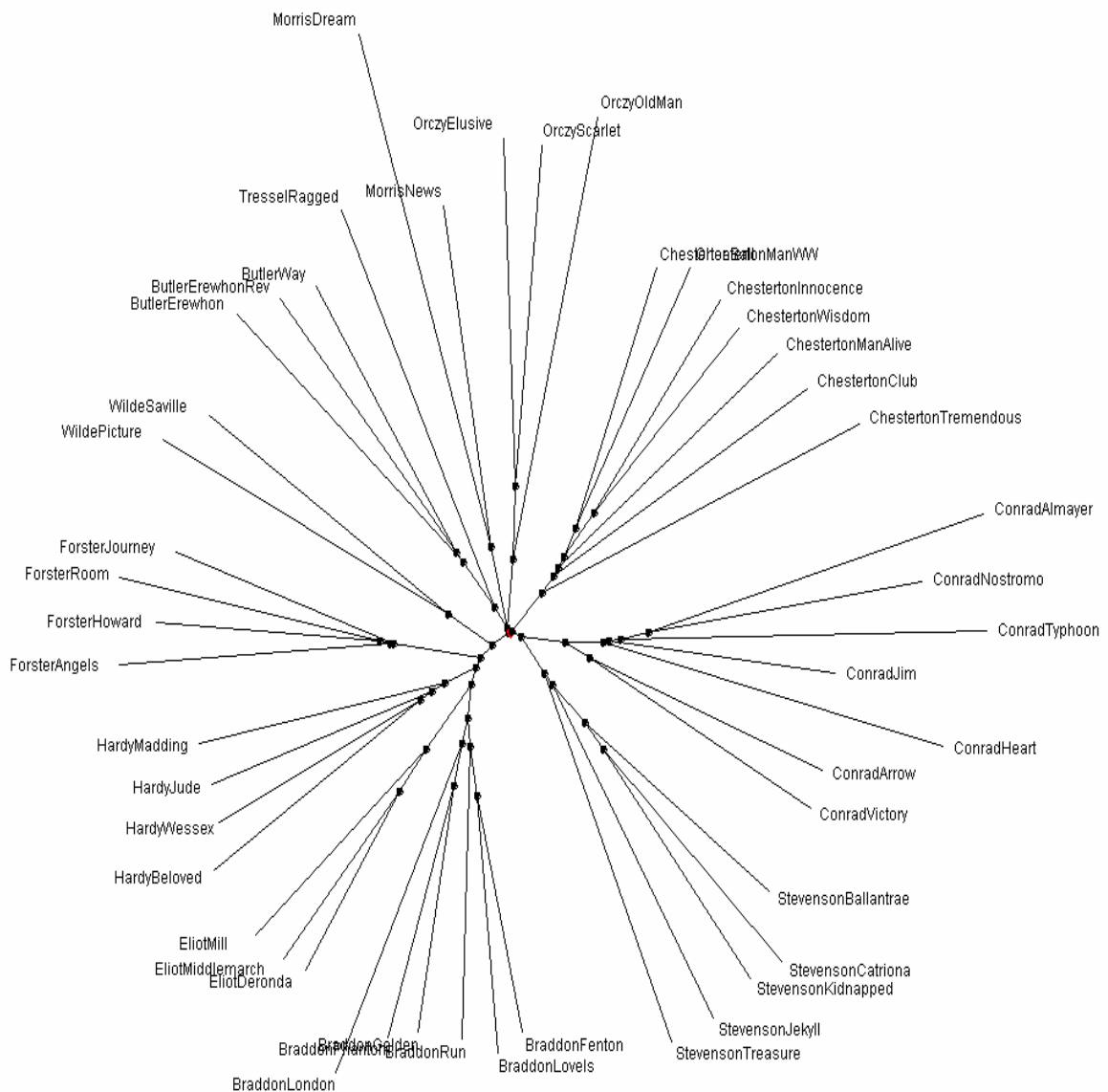
An agreement with Dr Ledger was reached about the main characteristics of this new corpus. It should contain the entire novels from which the Corpus Oxquarry1 was composed in order to make the results comparable. And it was decided to add other novels written during the same period (1870-1914) by other writers. The main purpose was to obtain, by a choice independent of the operator, the largest possible range of sizes for texts written in the same genre, period and by undisputed authors.

The result of this selection is Corpus Oxquarry2 (Appendix 3). This corpus is 5.1 million tokens long (ten times Oxquarry1) with an average length slightly over 100,000 tokens. The length ratio is nearly 1:16 between Heart of Darkness by Joseph Conrad (20,243 tokens) and Middlemarch by George Eliot (316,156 tokens). This ratio of lengths clearly exceeds the limits of validity of formula (3) observed by Labbé and Labbé (2003). Beyond a scale of 1:10, the experiments conducted on French corpora demonstrate that distance indices are biased, especially when calculations include the longest and the shortest texts.

This bias should not be over-emphasized. For example, tree-classification — applied to the whole matrix of indices calculated on Corpus Oxquarry2 applying Formula 3 — is without error (Fig. 5 below). Of course, the lengths of the paths must be carefully considered. For example, for Conrad's cluster, some novels appear to be very near one another (Lord Jim, Arrow of Gold or Victory), while others are more outlying (Almayer's Folly, Typhoon, Heart of Darkness). One must not draw conclusions too quickly about Conrad's work because the first set contains much longer texts than the second one. In fact, the first experiment above

allows one to "predict" that, the influence of the other factors being equal, the longer texts will be relatively nearer all the others than the shorter ones.

Fig. 5. Tree classification of Corpus Oxquarry2 applying Formula 3



Even if tree-classification appears to be accurate, one must bear in mind that this classification does not yield "proof" of authorship, but only provides the "best fit" of the whole population. Its suggestions must be confirmed by the direct examination of the matrix as in the blind-test presented above. Here, the mean distance index is equal to .315, greatly below the one obtained for Oxquarry1 (.405). In other words, when the mean sizes of the texts are multiplied by ten, their distance indices decrease by nearly 20% on average.

But this decrease is not linear and strictly proportional. There are also some irregularities which render the method presented above no longer reliable for authorship attribution. The

5% lower limit (mean less 1.96 standard deviations) is equal to .237; the upper limit to .394.

All pairs of novels separated by a distance greater than .394 are indeed by two different authors. By contrast, the lower boundary generates three errors of attribution (bold type in Table 9).

Table 9. Inter-textual distance indices calculated on Corpus Oxquarry2 applying formula 3 - Pairs of distance indices which suggest same authorship (mean less 1.96 standard deviations)

Text 1	Text 2	Indice
EliotDeronda	EliotMiddlemarch	0.155
BraddonFenton	BraddonLovels	0.180
BraddonGolden	BraddonPhantom	0.185
ChestertonInnocence	ChestertonWisdom	0.186
EliotMiddlemarch	EliotMill	0.196
EliotDeronda	EliotMill	0.196
ForsterHoward	ForsterJourney	0.202
BraddonFenton	BraddonRun	0.204
ConradJim	ConradVictory	0.204
BraddonGolden	BraddonLovels	0.205
HardyJude	HardyWessex	0.206
ChestertonBall	ChestertonManWW	0.206
HardyBeloved	HardyWessex	0.207
HardyMadding	HardyWessex	0.208
BraddonLovels	BraddonPhantom	0.211
ForsterHoward	ForsterRoom	0.214
StevensonCatriona	StevensonKidnapped	0.214
HardyBeloved	HardyJude	0.216
StevensonBallantrae	StevensonKidnapped	0.219
HardyJude	HardyMadding	0.219
ChestertonBall	ChestertonInnocence	0.219
ForsterAngels	ForsterJourney	0.219
ConradArrow	ConradVictory	0.220
StevensonBallantrae	StevensonCatriona	0.221
BraddonFenton	BraddonGolden	0.222
BraddonLovels	BraddonRun	0.222
ForsterJourney	ForsterRoom	0.224
ChestertonInnocence	ChestertonManWW	0.226
ForsterAngels	ForsterHoward	0.228
ChestertonBall	ChestertonWisdom	0.229
BraddonLondon	BraddonPhantom	0.230
BraddonLovels	EliotDeronda	0.230
ForsterAngels	ForsterRoom	0.231
BraddonFenton	EliotMiddlemarch	0.231
ButlerErewhonRev	ButlerWay	0.232
ChestertonBall	ChestertonManAlive	0.233
EliotMill	HardyJude	0.234
BraddonFenton	BraddonPhantom	0.234
BraddonLovels	EliotMiddlemarch	0.235
BraddonGolden	BraddonRun	0.236
BraddonPhantom	BraddonRun	0.237
ChestertonManWW	ChestertonWisdom	0.237

As predicted from the discussion above, these errors involve the longest texts. Indeed, it is possible that the four novels and the two authors (Braddon and Eliot) were in fact "neighbours" as to their content and vocabularies, but also it is clear that Formula (3) — when applied to texts the size of which differs greatly (here more than 1:15) — "advantages" the longer ones, ie: the longer ones appear to be nearer one another — and also from all the others — than the shorter ones.

In order to measure the strength of association between distance index and length of the texts examined, a calculation is proposed. The comparison of each text with the others generates 47 indices. If the Formula 3 "advantages" the longer texts, the averages of their 47 indices must be lower than the same averages calculated for the shorter ones. The results of all these calculations are summarised in Fig. 6.

Fig. 6. Correlation between the length of each text and its average distance index compared with all the other texts (Corpus Oxquarry2).

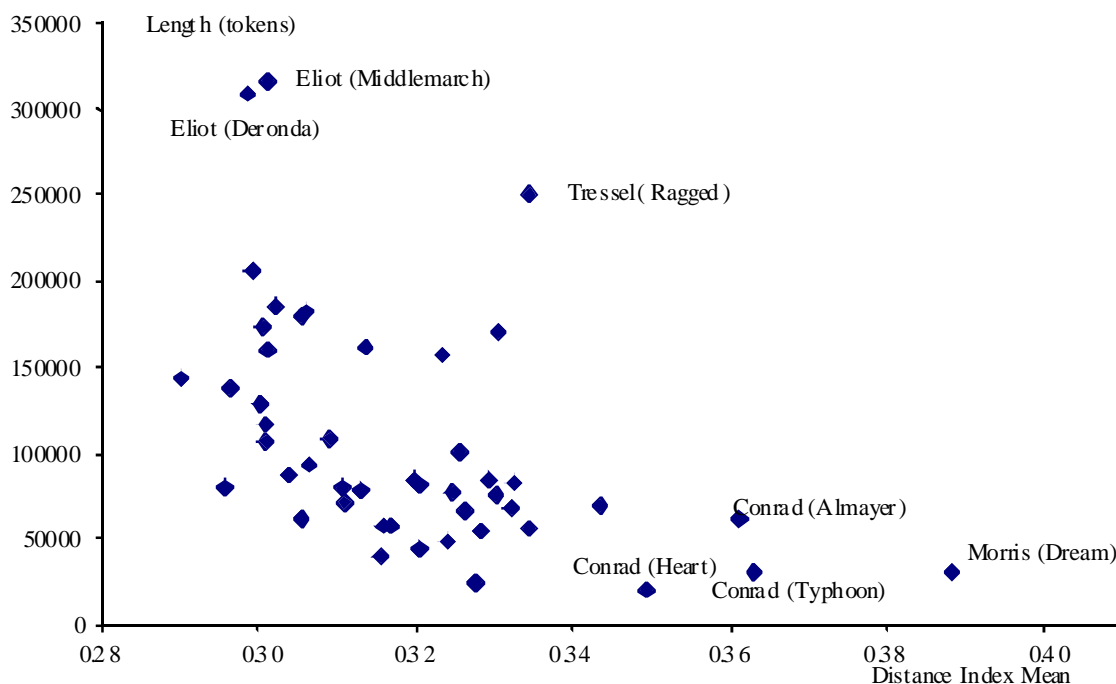


Fig. 6 suggests an inverse relation between lengths of texts and distance indices. The plotted points are oriented NW-SE and present a more or less hyperbolic shape: the shorter the texts, the farther apart they appear to be in relation with the others. But it is noticeable that 7 texts stand out apart from the others. On the lower right, four of the shortest texts: Dream of J.B by Morris and three novels by Conrad (Typhoon, Almayer's Folly, Heart of Darkness); on the upper left, the three longer texts: Tressel's Ragged Philantropists, Eliot's Middlemarch and

Daniel Deronda. When these 7 texts are removed, there is no longer a significant negative correlation between length of each text and the averaged distance indices. Of course the "size factor" is already present in these data, but it is "masked" by stronger factors: authorship and content. It is note-worthy that, once these 7 texts are extracted, the ratio between the shortest remaining text (Stevenson's Doctor Jekyll) and the longest remaining one (Braddon's The Mill) is equal to 1:8. Thus it is possible to draw a provisional conclusion: the procedure of authorship attribution, presented above in the blind-test, should be used only when the lengths of the texts compared are within a ratio less than 1:8.

2.4. Second large scale experiment

In order to measure accurately the effect of text length on the distance index, an experiment was conducted on Corpus Oxquarry2. A large number (100) of slices of equal size (n contiguous tokens) was drawn at random from each text of Corpus Oxquarry2 and formula (1) was applied to all the different pairs formed with all the equivalent slices drawn from all the other texts. Thus, the distance index between a pair of texts was the mean of 10,000 different calculations (for the reasons for this choice, see below, section 2.6). The experiment started with an excerpt size of 5,000 tokens; then this length was increased by 1,000 tokens and so on... until 20,000 tokens (summary is given in Table 10). It was not possible to go further than 20,000 tokens because of the length of the shortest novel (Conrad's Heart of Darkness: 20,243 tokens). During the experiment, the minimum index was between Chesterton's The Innocence of Father Brown and The Wisdom of Father Brown (identity of authorship and proximity of the two topics). The maximum index was between Butler's Erewhon and Conrad's Almayer's Folly which is also not surprising because of the wide disparity of topics which adds its effect to the expected effect of differing authorship.

Table 10 Main results of a large scale experiment on Corpus Oxquarry2

Excerpt sizes (tokens)	Mean Distance Index	Minimum Distance Index	Maximum Distance Index	Standard deviation	Variation coefficient
5 000	0.4578	0.3775	0.5474	0.0244	0.0532
6 000	0.4463	0.3657	0.5428	0.0256	0.0574
7 000	0.4365	0.3547	0.5361	0.0269	0.0617
8 000	0.4277	0.3459	0.5305	0.0271	0.0634
9 000	0.4207	0.3372	0.5265	0.0279	0.0664
10 000	0.4137	0.3300	0.5203	0.0279	0.0674
11 000	0.4090	0.3240	0.5200	0.0290	0.0710
12 000	0.4038	0.3166	0.5161	0.0293	0.0726
13 000	0.3996	0.3097	0.5135	0.0300	0.0750
14 000	0.3954	0.3046	0.5106	0.0303	0.0766
15 000	0.3916	0.2991	0.5075	0.0304	0.0777
16 000	0.3882	0.2941	0.5051	0.0309	0.0797
17 000	0.3850	0.2889	0.5031	0.0311	0.0809
18 000	0.3821	0.2845	0.5016	0.0316	0.0828
19 000	0.3794	0.2807	0.5002	0.0318	0.0838
20 000	0.3766	0.2768	0.4978	0.0320	0.0849

*

The index mean slowly decreases as the length of the excerpts increases. As shown on Fig. 7 below, this movement is not linear but more or less hyperbolic. The slight irregularities in the curves are usual in any random process, even if the number of excerpts is large. One must also note that the origin of Y-axis is not placed at zero but at .25 in order to highlight the fluctuations of the indices: the distance indices decrease slowly and the rate of decrease is becomes less as the size of excerpts increases. Conclusions of the first experiment, on excerpts by Stevenson and Orczy, are verified. It is also clear that the influence of the "size factor" is more acutely felt on short texts - especially those under 10,000 tokens - than on longer ones.

Another characteristic is highlighted by the graph: the range between minimum and maximum indices widens as excerpt size expands. This change does not derive from a small number of marginal texts; it applies to all the series as indicated by the evolution of the standard deviation. If the variation of excerpt size did not influence the dispersion of the distance indices around the mean, the "coefficient of variation" (ratio of standard deviation/mean) would be stable. On the contrary, this coefficient increases regularly (Fig. 8). In other words, when the scale of observation is reduced, differences between texts are slightly but significantly eroded.

Fig. 7 Evolution of the Index Mean and the Min/Max Range in relation to size of excerpts

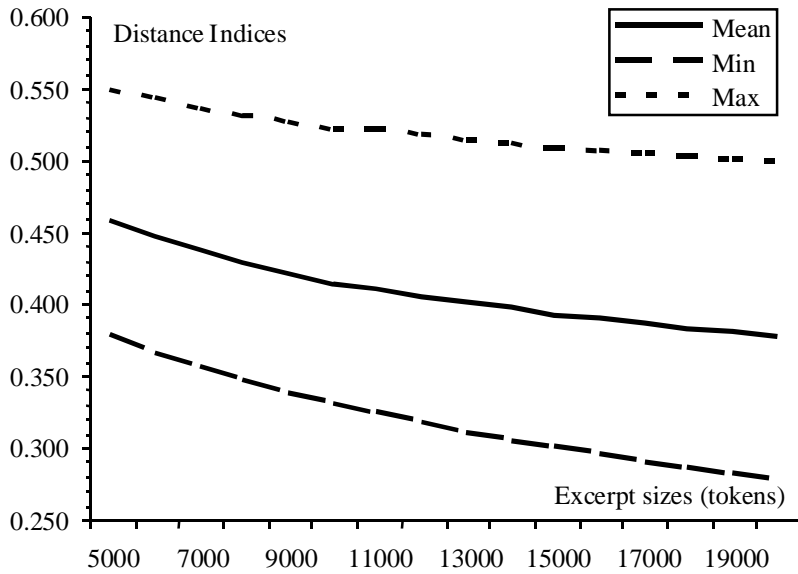
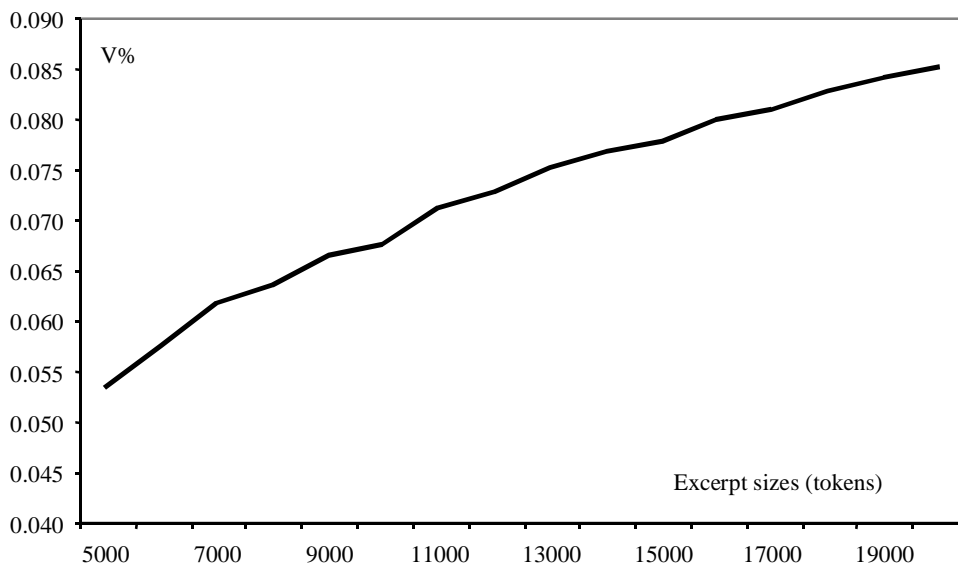


Fig. 8 Evolution of the ratio, Standard Deviation/Mean, related to the size of excerpts



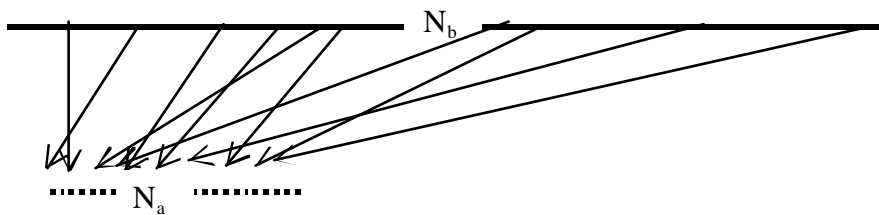
The main explanation is as follows: as the size of a text increases, the relative weight of the low frequency words decreases with respect to its whole vocabulary. Since, as shown above, the low frequency words contribute most to the distance index, the values of this index logically decrease as the text lengths increase. For French texts, lemmatisation greatly reduces the weight of these low frequency words. As a result, the influence of the length on the distance index is distinctly lower than with English texts.

2.5 Neutralisation of the length factor: the "sliding window method"

How to neutralise the influence of the length factor in order to conduct authorship attribution experiments free from bias?

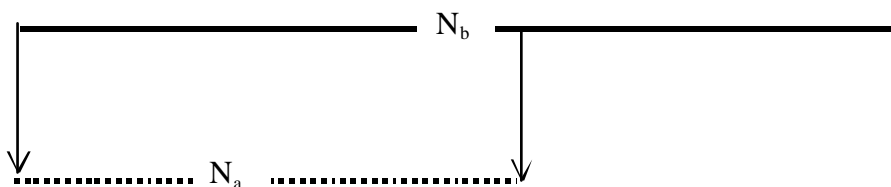
Formula (3) can be described as a process of drawing N_a word types throughout the whole text B , B been greater than A (Fig. 9 below).

Fig. 9 How Formula 3 operates



Applied to texts which were written in the same genre and period, this method usually "advantages" the longer text (B) for reasons shown above. To obviate this drawback, Tom Merriam (2002b et 2003a) proposed to consider only one block of the N_a first contiguous words drawn out of text B (Fig. 10).

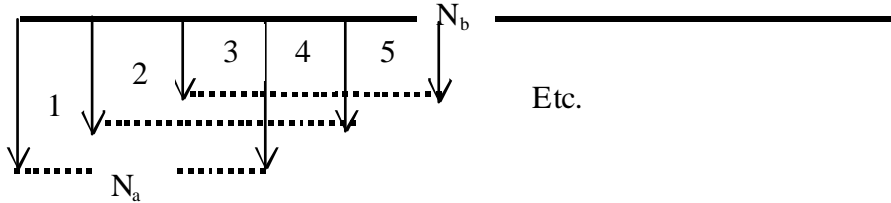
Fig. 10 Tom Merriam's method



In T. Merriam's experiment, N_a was equal or superior to the half of the lengths of the whole texts considered. Admittedly, he acknowledges that : "Problems might arise in the case of a [text] with a first half written by one author and a second by another" (Merriam, 2003a). These problems appear to have a simple solution: draw by chance out of B a large number of excerpts of N_a contiguous words and, then, calculate the means of all the distance indices between all these excerpts and the text A . This solutions entails several drawbacks. First, an error margin should be associated with the indices and this may often confound any clear conclusion. Above all, some interesting properties of Euclidian distances are endangered, like identity, symmetry or triangle inequality.

In order to avoid these drawbacks, it is necessary that the calculation gives equal weight to each token in the two texts compared. With this aim in mind, it is proposed to use a "sliding window". For example, in the Fig. 11 below, the steps (or intervals) of the sliding windows are equal to $0.33 N_a$: each token of B is included three times in the calculation of the index.

Fig. 11 The "sliding window" method



Actually, it is possible to choose a standardised length (L) for these windows (or "blocks") of contiguous tokens. In which case, the distance index between two texts is the mean of all indices calculated with Formula (1) on all the excerpts (of equal size) drawn from each of the two texts.

Are the results reliable when this method is applied to texts whose sizes greatly differ? Let:

- A and B be two texts the lengths of which are N_a and N_b tokens.
- A and B be divided into M_a and M_b blocks of equal lengths (L contiguous tokens).
- their division be done along the two texts following the same interval of P tokens (see Figure 11 above), L being a multiple of P .

Formula (1) can be applied without bias to any pair of blocks, one of them (i) drawn from A and the other block (j) drawn from B — the length of these two blocks being equal to L contiguous tokens— all the properties of an Euclidian distance are preserved, especially the ones of identity, symmetry and triangle inequality. The inter-textual distance index between A and B becomes the mean of all the different indices of distance which can be calculated between all differing blocks drawn from these two texts, ie $M_a.M_b$:

$$(4) \quad D_{L(a,b)} = \frac{1}{M_a.M_b} \sum_{i=1}^{i=M_a} \sum_{j=1}^{j=M_b} D_{(a_i,b_j)} \quad \text{with } D_{(a_i,b_j)} \text{ calculated with (1)}$$

In order to measure its accuracy, this procedure and formula were applied in a large number of experiments carried out on Corpus Oxquarry1 and 2.

2.6 Accuracy of the "sliding window" method.

Corpus Oxquarry1 (made of texts the length of which is 10,000 tokens) was used in order to compare the results of the "sliding window" to those of the blind-test. Table 11 below presents a synthesis of the results of the experiments on blocks of 5,000 contiguous tokens, while varying the step interval.

Table 11

Synthesis of the results of experiment on Oxquarry 1, blocks of 5,000 contiguous tokens.

Step interval in tokens	1,000	500	200
Number of calculations (Ma.Mb)	100	400	2,500
Deviations between observed and expected values:			
Absolute mean	0.62%	0.59%	0.61%
Standard deviation	0.52%	0.49%	0.49%
Margin of error	± 1.66%	± 1.56%	± 1.59%

When a step interval of 200 tokens is chosen (as in the last column of Table 11), it means that the first window starts at the first token and ends at the 5,000th token, the second starts at 201 and finishes at 5,200 and so on... until the last one which starts at 9,801 and finishes at 4,800 (the algorithm "buckles" in order to form the last blocks). Thus, each word of *A* and *B* appears in 50 different blocks of 5,000 tokens, and the distance index between each pair of texts (*A*, *B*) is the mean of 2,500 calculations (each of the 50 blocks of *A* is compared to each of the 50 blocks of *B*).

The results in Table 11 show that, in superimposing the two 1,326 cell matrices, the average deviation between expected and observed indices is appreciably less than 1% and that more than 19 twentieths of these discrepancies are contained within a fluctuation margin of ±1.6% about this mean. In fact, 38 values are above or below these two thresholds. As predicted above, these anomalies derive from the nearest texts. For example, the greatest deviation (-3.1%) was observed between 1H and 2Z (both excerpts from Far the madding crowd by Hardy); the next in order of magnitude (+2.8% and + 2.6%) are from 4 excerpts from Catrinae by Stevenson. Two explanations can be offered. Firstly the experiments, presented in the first section of this paper, show that, in case of a great proximity in the vocabulary of the two texts, the augmentation of *L* makes their intertextual distance decrease faster than the average, and conversely, when the two texts are farther apart (see also Fig. 6

above). Stevenson's texts illustrate one case and Orczy's novels the other. Secondly, the buckle method adds a slight but real unaccounted - for variability.

Table 11, first line, shows the number of calculations entailed by the size of the step interval: this number grows exponentially as the step size is reduced. So a best compromise should be decided on, taking into consideration the demands of accuracy and the necessity of not waiting hours for results. Table 11 suggests that the step should be chosen in order to obtain at least 100 different calculations, and that no further accuracy can be expected in using over 400 different calculations.

Are these conclusions confirmed when applying the method to longer texts? Corpus Oxquarry2 was used in order to provide some answers.

Granted the fact that, in this corpus, the shortest text is slightly over 20,000 tokens, it was decided to adopt this size for L and a P (the size of the step) of 2,000 in order to ensure that the number of calculations between two texts are always greater than 100. The average distance index is .3712 and the standard deviation .0282. So it can be supposed that all the novels which are separated by indices less than .315 (the mean less 1.96 standard deviations) are by the same authors. Table 12 below shows that this authorship attribution is free of error — contrary to Table 9 — but also that there is no text by Morris or by Wilde. One must bear in mind that, at this stage, the aim was not to give an answer for all the texts, but to avoid error.

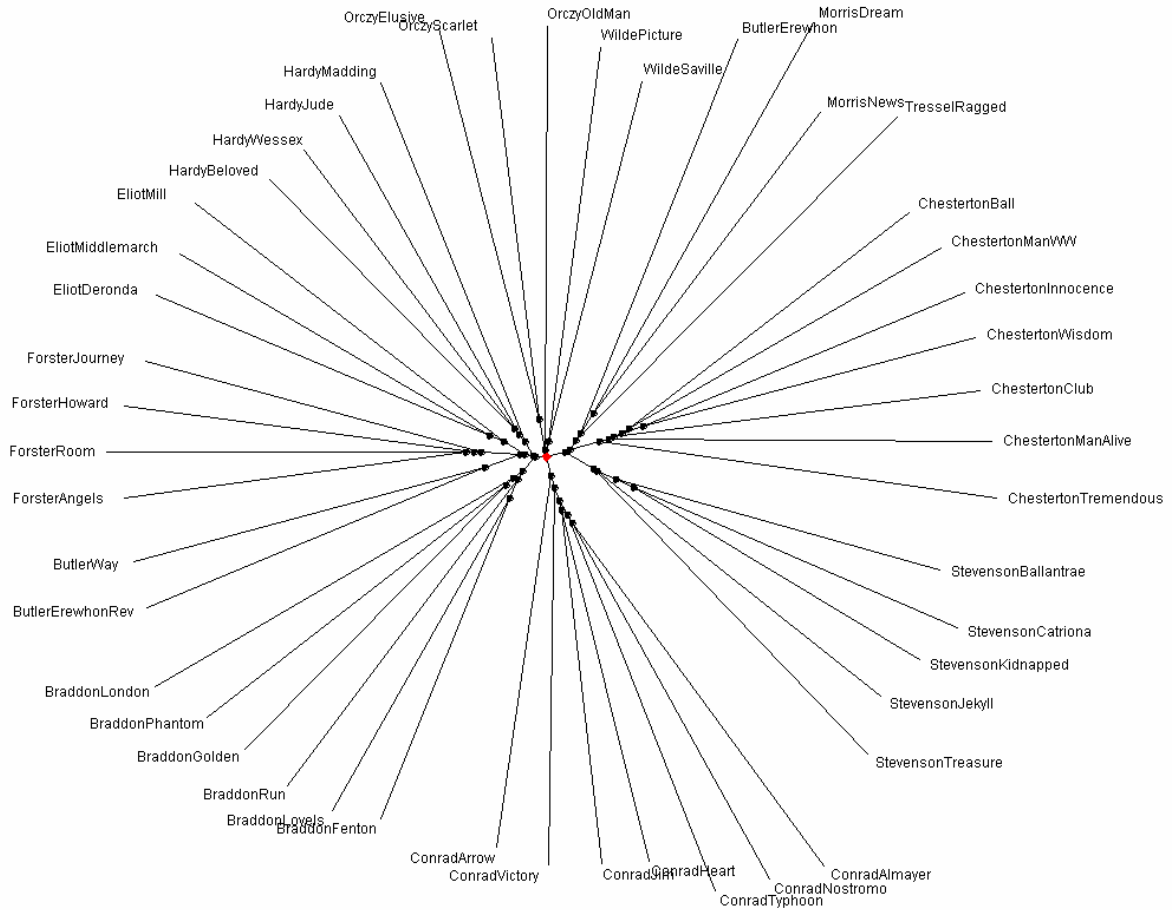
Table 12. Corpus Oxquarry2. Pairs of texts separated by an index lower than .315
(blocks of 20,000 tokens, step of 2,000 tokens).

Text 1	Text 2	Indices
ChestertonInnocence	ChestertonWisdom	0.277
HardyBeloved	HardyWessex	0.282
ForsterAngels	ForsterJourney	0.283
StevensonCatriona	StevensonKidnapped	0.286
ChestertonBall	ChestertonManWW	0.287
ForsterAngels	ForsterRoom	0.289
StevensonBallantrae	StevensonJekyll	0.289
StevensonBallantrae	StevensonKidnapped	0.290
ForsterHoward	ForsterJourney	0.292
ConradHeart	ConradJim	0.292
BraddonFenton	BraddonLovels	0.292
ForsterAngels	ForsterHoward	0.292
ForsterHoward	ForsterRoom	0.292
EliotDeronda	EliotMiddlemarch	0.295
StevensonBallantrae	StevensonCatriona	0.295
HardyBeloved	HardyJude	0.297
ChestertonBall	ChestertonInnocence	0.299
ForsterJourney	ForsterRoom	0.299
ChestertonInnocence	ChestertonManWW	0.299
StevensonKidnapped	StevensonTreasure	0.300
HardyJude	HardyWessex	0.301
HardyMadding	HardyWessex	0.302
ChestertonClub	ChestertonManWW	0.303
StevensonJekyll	StevensonKidnapped	0.303
ChestertonBall	ChestertonWisdom	0.304
ChestertonBall	ChestertonClub	0.305
ChestertonManWW	ChestertonWisdom	0.306
ConradJim	ConradVictory	0.306
ChestertonBall	ChestertonManAlive	0.307
ButlerErewhonRev	ButlerWay	0.308
BraddonFenton	BraddonRun	0.310
ConradArrow	ConradVictory	0.310
BraddonGolden	BraddonPhantom	0.310
ChestertonClub	ChestertonWisdom	0.312
ChestertonManAlive	ChestertonWisdom	0.312
ConradJim	ConradTyphoon	0.313
ChestertonClub	ChestertonInnocence	0.313
ChestertonClub	ChestertonManAlive	0.314
BraddonGolden	BraddonLovels	0.315

As was remarked about the blind-test with Oxquarry1, all the reasoning is based on two assumptions: (1) the corpus contains several authors and (2) there are, at least, two novels by each of these authors (this second assumption is not entirely true: there is only one book by Tressel).

Does tree-classification give a more complete and reliable answer? (Fig. 12)

Fig. 12 Tree-classification of Corpus Oxquarry2 using Formula (4) (window size: 20,000 tokens; step: 2,000 tokens).



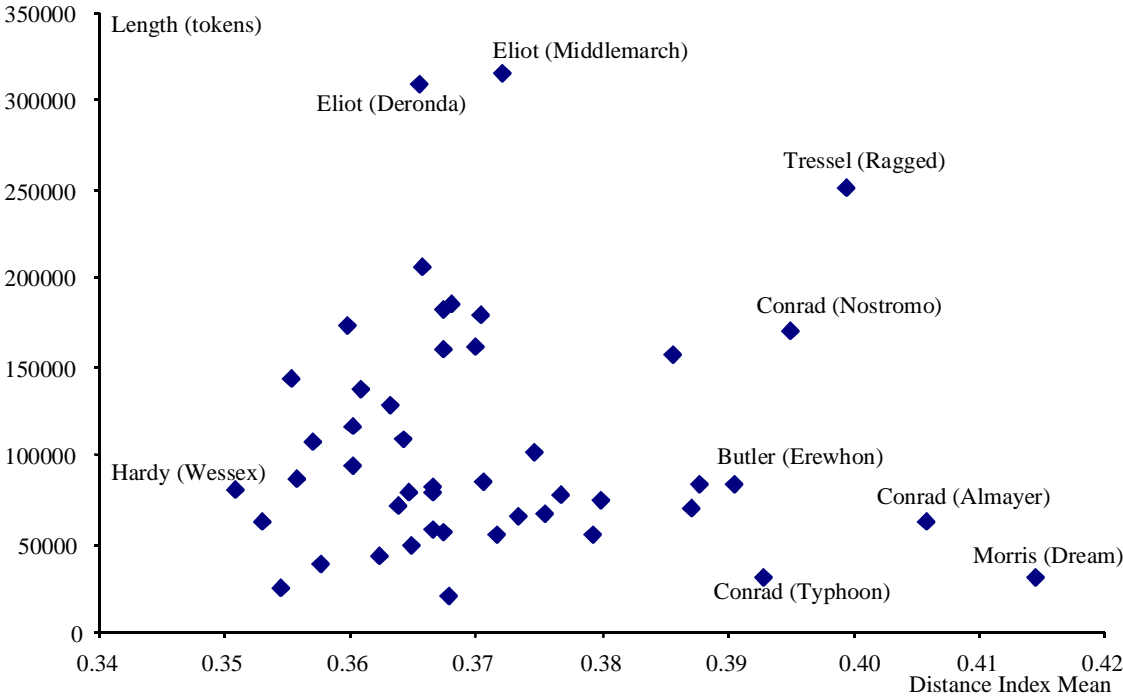
The confidence index for the whole tree is .97; for all the nodes, this index is greater than .90 and all the paths are over .80. It is thus possible to believe that the whole tree reproduces, almost faithfully, the information provided by the original matrix of distance indices, but that some of the terminal branches should be considered with care. The "problematic" texts (located in N-NE part of the tree) are also those which are the farthest from the others: Butler's Erewhon, Orczy's Old Man, Tressel's Ragged Philanthropists, or those which are relatively far from their "brothers" (other texts by the same author), and relatively near several others (like Wilde's two novels). None of these problematic texts appears in Table 12 and they are all joined to the root of the tree or very near it. In other words, the algorithm fails to classify them. Thus tree-classification must be improved by excluding these problematic texts from the tree.

In addition, the "star-shape" must also be noted. As G. Ledger supposed, the vocabularies of the majority of these novels are not very different from each other, even if they are by different authors. One must also admit that the sliding window method erodes, slightly but significantly, the differences between texts (when used on texts the lengths of which greatly differ).

But it must be borne in mind that all these drawbacks were predictable given the fact that the ratio of text lengths in Oxquarry2 was clearly beyond the valid limits of formula (3) and that this experiment was arranged precisely for this reason.

Given these limitations, this method really neutralises the effect of length differences on the distance index as can be observed in Fig. 13 (to be compared with the Fig. 6)

Fig. 13 Correlation between the length of each text and its average distance index with all the others ("sliding window" method: $L = 20,000$; $P = 2,000$).



It must also be noted that some books stand apart (on the X-axis): one by Morris (Dream...), the only book of Tressel, three by Conrad (Almayer, Nostromo, Typhoon), and one by Butler (Erewhon). Are these books different from the others because of their topics, of their style? Intertextual distance leads to an interesting question and offers a first possible response with the help of the "contribution to the distance"...

Conclusions

A necessary pre-condition must be kept in mind. The spelling of all words must be carefully checked, and a simple rule must be applied to the whole corpus: "one word, one spelling". This standardisation is necessary in order to be sure that the differences are not partly "artificial"...

If texts compared are not too different in size, Formula (3) can be used. When that is not the case, or when more accuracy is needed, the "sliding window" method provides a useful tool. Further experiments carried out on other texts demonstrate that the "sliding window" is not generally sensitive to differences of text size and that its results can be linked to a standardised scale of authorship attribution, like the one presented in the first section of this paper (for $L = 10,000$ tokens). In any case, it is preferable not to use these methods when differences of sizes exceed a ratio of 1:8, and to be very cautious when applying them on texts whose lengths are under 10,000 tokens.

Even if the notions of "genre", "topics" and "period" are to be more precisely discussed, these experiments show clearly that, for texts written in the same "genre" by contemporaneous writers, authorship is almost always the dominant factor.

In addition, the sliding window method offers a useful tool with which to analyse "collaborative" texts. When a text is known (or) suspected to have been written by several hands (e.g. Merriam, 2005), which are the parts by one or other(s) of the potential authors?

Formula (3) and (4) are easy to compute and some interesting characteristics of the two methods may be pointed out:

— Whole texts are considered and, in the calculations, each word type renders exactly its frequency weighting in the text without any "pre-selection" or presupposition about "characteristic" words of each author;

— There is also no prejudging of "sure" or "doubtful" authorship. Given the fact that the method entails no limitation on the number of texts used, it is possible to take into consideration all the available ones (on electronic files), subject to the restrictions of spelling standardisation and of scale ratio cited above.

These two points are important, given the "state of the art" in authorship attribution studies and the numerous critiques which are addressed to them (for example: Rudman, 1998).

Finally, the combination of intertextual distance with modern classification methods considerably furthers the single question of authorship attribution. One is now able to answer some very interesting questions. For example: in a given genre and period, which authors are

"neighbours" and which are distant? Which are the words which generated this proximity or opposition? Here lies a possible answer to Craig's question: "If you can tell authors apart, have you learned anything about them?" (Craig, 1999)...

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Appendix 1 "Corpus Oxquarry1" (all excerpt are 10,000 tokens long)

Except	Set 1 Author	Titles	Chap.	Except	Set 2 Author	Titles	Chap.
A	Hardy	Jude	I	A	Butler	Erewhon revisit.	XIV
B	Butler	Erewhon revisit.	II	B	Morris	Dream of JB	
C	Morris	News	XIII	C	Tressel	Ragged TP	
D	Stevenson	Catrinae	V	D	Hardy	Jude	
E	Butler	Erewhon revisit.	XVIII	E	Stevenson	Ballantrae	IV
F	Stevenson	Ballantrae	II	F	Hardy	Wessex Tales	
G	Conrad	Lord Jim	XIV	G	Orczy	Elusive P	VII
H	Hardy	Madding	III	H	Conrad	Lord Jim	XXI
I	Orczy	Scarlet P	I	I	Morris	News	VIII
J	Morris	Dream of JB	VII	J	Hardy	Well beloved	I
K	Stevenson	Catrinae	X	K	Conrad	Almayer	VI
L	Hardy	Jude	VII	L	Hardy	Well beloved	XII
M	Orczy	Scarlet P	XIV	M	Morris	News	XIX
N	Stevenson	Ballantrae	V	N	Conrad	Almayer	XI
O	Conrad	Lord Jim	VII	O	Forster	Room with view	I
P	Chesterton	Man who was	I	P	Forster	Room with view	IV
Q	Butler	Erewhon revisit.	VII	Q	Conrad	Almayer	IX
R	Chesterton	Man who was	VII	R	Stevenson	Catrinae	XVI
S	Morris	News	I	S	Hardy	Madding	X
T	Conrad	Almayer	II	T	Hardy	Well beloved	2 VI
U	Orczy	Elusive P	I	U	Chesterton	Man who was	III
V	Conrad	Lord Jim	II	V	Forster	Room with view	VIII
W	Orczy	Elusive P	XIV	W	Stevenson	Catrinae	I
X	Hardy	Wessex Tales		X	Hardy	Well beloved	VIII
Y	Tressel	Ragged TP		Y	Orczy	Scarlet P	VII
Z	Tressel	Ragged TP		Z	Hardy	Madding	XVIII

Appendix 2.1

Lowest intertextual distance indices ranked by ascending values (Corpus Oxquarry1)

Rank	Pair of texts	Index
1	1D (StevensonCatri) - 2W (StevensonCatri)	0.259
2	1N (StevensonBall) - 2E (StevensonBall)	0.265
3	1F (StevensonBall) - 2E (StevensonBall)	0.267
4	1K (StevensonCatri) - 2W (StevensonCatri)	0.277
5	1S (MorrisNews) - 2M (MorrisNews)	0.281
6	2N (ConradAlma) - 2Q (ConradAlma)	0.283
7	2L (HardyWell) - 2T (HardyWell)	0.283
8	1C (MorrisNews) - 2I (MorrisNews)	0.288
9	2J (HardyWell) - 2L (HardyWell)	0.289
10	1D (StevensonCatri) - 1K (StevensonCatri)	0.291
11	1E (ButlerErew) - 1Q (ButlerErew)	0.291
12	2R (StevensonCatri) - 2W (StevensonCatri)	0.291
13	2J (HardyWell) - 2T (HardyWell)	0.293
14	1F (StevensonBall) - 1N (StevensonBall)	0.294
15	2I (MorrisNews) - 2M (MorrisNews)	0.297
16	1D (StevensonCatri) - 2R (StevensonCatri)	0.302
17	2O (ForsterRoom) - 2P (ForsterRoom)	0.302
18	1S (MorrisNews) - 2I (MorrisNews)	0.303
19	1B (ButlerErew) - 1Q (ButlerErew)	0.304
20	1Z (TresselRagg) - 2C (TresselRagg)	0.304
21	2J (HardyWell) - 2X (HardyWell)	0.306
22	1F (StevensonBall) - 2R (StevensonCatri)	0.307
23	2E (StevensonBall) - 2R (StevensonCatri)	0.309
24	1L (HardyJude) - 2J (HardyWell)	0.310
25	1B (ButlerErew) - 2A (ButlerErew)	0.310
26	1J (MorrisDream) - 2B (MorrisDream)	0.311
27	1K (StevensonCatri) - 2R (StevensonCatri)	0.311
28	1P (ChestertonMan) - 2U (ChestertonMan)	0.312
29	2L (HardyWell) - 2X (HardyWell)	0.312
30	2T (HardyWell) - 2X (HardyWell)	0.313
31	1B (ButlerErew) - 1E (ButlerErew)	0.314
32	1L (HardyJude) - 2D (HardyJude)	0.314
33	1L (HardyJude) - 2L (HardyWell)	0.314
34	1M (Scar) - 2Y (OrczyScar)	0.314
35	1X (HardyWess) - 2L (HardyWell)	0.314
36	1M (Scar) - 1W (OrczyElus)	0.315
37	2O (ForsterRoom) - 2V (ForsterRoom)	0.315
38	1O (ConradLordJ) - 2H (ConradLordJ)	0.317
39	1P (ChestertonMan) - 1R (ChestertonMan)	0.317
40	1G (ConradLordJ) - 1O (ConradLordJ)	0.318
41	1H (HardyMadd) - 2Z (HardyMadd)	0.318
42	1Q (ButlerErew) - 2A (ButlerErew)	0.318
43	1R (ChestertonMan) - 2U (ChestertonMan)	0.319

Rank	Pair of texts	Index
44	1D (StevensonCatri) - 1N (StevensonBall)	0.319
45	1D (StevensonCatri) - 2E (StevensonBall)	0.320
46	1X (HardyWess) - 2J (HardyWell)	0.320
47	1N (StevensonBall) - 2R (StevensonCatri)	0.320
48	2F (HardyWess) - 2T (HardyWell)	0.321
49	2D (HardyJude) - 2J (HardyWell)	0.322
50	2K (ConradAlma) - 2Q (ConradAlma)	0.323
51	2D (HardyJude) - 2L (HardyWell)	0.323
52	1H (HardyMadd) - 2S (HardyMadd)	0.325
53	1A (HardyJude) - 2D (HardyJude)	0.325
54	1W (OrczyElus) - 2G (OrczyElus)	0.325
55	1X (HardyWess) - 2T (HardyWell)	0.326
56	1F (StevensonBall) - 1K (StevensonCatri)	0.326
57	2S (HardyMadd) - 2Z (HardyMadd)	0.326
58	1X (HardyWess) - 2X (HardyWell)	0.327
59	1K (StevensonCatri) - 1N (StevensonBall)	0.327
60	2K (ConradAlma) - 2N (ConradAlma)	0.327
61	1L (HardyJude) - 2T (HardyWell)	0.329
62	2E (StevensonBall) - 2W (StevensonCatri)	0.329
63	1E (ButlerErew) - 2A (ButlerErew)	0.330
64	1L (HardyJude) - 1X (HardyWess)	0.330
65	2D (HardyJude) - 2T (HardyWell)	0.330
66	1G (ConradLordJ) - 2H (ConradLordJ)	0.330
67	1F (StevensonBall) - 2W (StevensonCatri)	0.330
68	1X (HardyWess) - 2F (HardyWess)	0.331
69	2F (HardyWess) - 2J (HardyWell)	0.331
70	1N (StevensonBall) - 2W (StevensonCatri)	0.332
71	1D (StevensonCatri) - 1F (StevensonBall)	0.333
72	1L (HardyJude) - 2F (HardyWess)	0.333
73	1H (HardyMadd) - 2J (HardyWell)	0.335
74	2D (HardyJude) - 2X (HardyWell)	0.335
75	1Y (TresselRagg) - 1Z (TresselRagg)	0.337
76	1H (HardyMadd) - 2T (HardyWell)	0.337
77	2F (HardyWess) - 2L (HardyWell)	0.338
78	1K (StevensonCatri) - 2E (StevensonBall)	0.338
79	2P (ForsterRoom) - 2V (ForsterRoom)	0.338
80	1X (HardyWess) - 2D (HardyJude)	0.339
81	1H (HardyMadd) - 2D (HardyJude)	0.340
82	1H (HardyMadd) - 2L (HardyWell)	0.340
83	1H (HardyMadd) - 1L (HardyJude)	0.344
84	1L (HardyJude) - 2X (HardyWell)	0.344
85	1W (OrczyElus) - 2Y (OrczyScar)	0.344
86	1K (StevensonCatri) - 1S (MorrisNews)	0.349

Appendix 2.2

Pairs of texts of which the inter-textual distance indices are between .35 and .36
(bold types : same authorship)

Rank	Pairs of texts	Indices
87	1S (MorrisNews) - 2W (StevensonCatri)	0.350
88	1A (HardyJude) - 1L (HardyJude)	0.350
89	2G (OrczyElus) - 2Y (OrczyScar)	0.350
90	1A (HardyJude) - 2J (HardyWell)	0.350
91	1Y (TresselRagg) - 2C (TresselRagg)	0.351
92	1F (StevensonBall) - 1G (ConradLordJ)	0.351
93	1H (HardyMadd) - 2F (HardyWess)	0.351
94	1C (MorrisNews) - 2M (MorrisNews)	0.351
95	2J (HardyWell) - 2Z (HardyMadd)	0.351
96	1M (OrczyScar) - 2G (OrczyElus)	0.352
97	1A (HardyJude) - 1H (HardyMadd)	0.353
98	1C (MorrisNews) - 1S (MorrisNews)	0.353
99	1L (HardyJude) - 1Z (TresselRagg)	0.353
100	1L (HardyJude) - 2S (HardyMadd)	0.354
101	2F (HardyWess) - 2X (HardyWell)	0.354
102	1A (HardyJude) - 2T (HardyWell)	0.355
103	1Q (ButlerErew) - 2J (HardyWell)	0.355
104	2T (HardyWell) - 2Z (HardyMadd)	0.355
105	1H (HardyMadd) - 1X (HardyWess)	0.355
106	1G (ConradLordJ) - 1N (StevensonBall)	0.356
107	1H (HardyMadd) - 2X (HardyWell)	0.357
108	2J (HardyWell) - 2S (HardyMadd)	0.357
109	1G (ConradLordJ) - 2E (StevensonBall)	0.357
110	1O (ConradLordJ) - 1V (ConradLordJ)	0.357
111	1Q (ButlerErew) - 2L (HardyWell)	0.357
112	1S (MorrisNews) - 2R (StevensonCatri)	0.357
113	1T (ConradAlma) - 2K (ConradAlma)	0.357
114	1F (StevensonBall) - 1S (MorrisNews)	0.357
115	2D (HardyJude) - 2Z (HardyMadd)	0.358
116	1F (StevensonBall) - 2H (ConradLordJ)	0.358
117	1A (HardyJude) - 2L (HardyWell)	0.358
118	2D (HardyJude) - 2H (ConradLordJ)	0.359
119	1A (HardyJude) - 1X (HardyWess)	0.359
120	2H (ConradLordJ) - 2X (HardyWell)	0.359
121	1X (HardyWess) - 2Z (HardyMadd)	0.359
122	2D (HardyJude) - 2F (HardyWess)	0.359
123	2D (HardyJude) - 2S (HardyMadd)	0.360
124	2F (HardyWess) - 2S (HardyMadd)	0.360
125	1L (HardyJude) - 1Q (ButlerErew)	0.360
126	1A (HardyJude) - 2X (HardyWell)	0.360
127	2L (HardyWell) - 2Z (HardyMadd)	0.360

Appendix 2.3

Pairs of excerpts by a same author separated by intertextual distance indices over .36

Rank	Pairs of texts	Indices
134	2F (HardyWess) - 2Z (HardyMadd)	0.363
139	1I (OrczyScar) - 2Y (OrczyScar)	0.364
143	2L (HardyWell) - 2S (HardyMadd)	0.365
144	1L (HardyJude) - 2Z (HardyMadd)	0.365
145	2S (HardyMadd) - 2T (HardyWell)	0.365
151	1X (HardyWess) - 2S (HardyMadd)	0.366
153	1G (ConradLordJ) - 1V (ConradLordJ)	0.367
157	1A (HardyJude) - 2F (HardyWess)	0.367
171	1U (OrczyElus) - 2Y (OrczyScar)	0.370
172	2X (HardyWell) - 2Z (HardyMadd)	0.370
179	1J (MorrisDream) - 2I (MorrisNews)	0.371
186	1V (ConradLordJ) - 2H (ConradLordJ)	0.373
187	1F (StevensonBall) - 1R (ChestertonManW)	0.373
209	1U (OrczyElus) - 1W (OrczyElus)	0.376
213	1A (HardyJude) - 2Z (HardyMadd)	0.376
224	1A (HardyJude) - 2S (HardyMadd)	0.378
229	1S (MorrisNews) - 2B (MorrisDream)	0.378
233	1C (MorrisNews) - 1J (MorrisDream)	0.379
235	1V (ConradLordJ) - 2N (ConradAlma)	0.379
253	1J (MorrisDream) - 2M (MorrisNews)	0.380
256	1I (OrczyScar) - 1U (OrczyElus)	0.381
261	1J (MorrisDream) - 1S (MorrisNews)	0.381
265	1U (OrczyElus) - 2G (OrczyElus)	0.381
274	2H (ConradLordJ) - 2N (ConradAlma)	0.382
305	1T (ConradAlma) - 2Q (ConradAlma)	0.384
307	1O (ConradLordJ) - 2N (ConradAlma)	0.385
343	1O (ConradLordJ) - 2Q (ConradAlma)	0.387
365	1V (ConradLordJ) - 2K (ConradAlma)	0.389
408	2S (HardyMadd) - 2X (HardyWell)	0.391
420	1T (ConradAlma) - 2N (ConradAlma)	0.392
463	1M (OrczyScar) - 1U (OrczyElus)	0.395

The remaining 863 indices (over .395) are all between excerpts by different authors

Appendix 2.4
 Pairs of excerpts separated by intertextual distance indices over .48

1301	1N (StevensonBall) - 1T (ConradAlma)	0.481
1302	2K (ConradAlma) - 2V (ForsterRoom)	0.483
1303	1D (StevensonCatri) - 1V (ConradLordJ)	0.483
1304	1T (ConradAlma) - 1Z (TresselRagg)	0.484
1305	1T (ConradAlma) - 2R (StevensonCatri)	0.486
1306	2K (ConradAlma) - 2W (StevensonCatri)	0.486
1307	1D (StevensonCatri) - 2K (ConradAlma)	0.486
1308	1J (MorrisDream) - 2K (ConradAlma)	0.488
1309	1J (MorrisDream) - 1V (ConradLordJ)	0.489
1310	1T (ConradAlma) - 2M (MorrisNews)	0.490
1311	1T (ConradAlma) - 2O (ForsterRoom)	0.490
1312	1C (MorrisNews) - 1T (ConradAlma)	0.491
1313	1T (ConradAlma) - 2A (ButlerErew)	0.491
1314	1T (ConradAlma) - 2B (MorrisDream)	0.492
1315	1T (ConradAlma) - 2C (TresselRagg)	0.494
1316	1T (ConradAlma) - 2I (MorrisNews)	0.496
1317	1S (MorrisNews) - 1T (ConradAlma)	0.496
1318	1B (ButlerErew) - 1T (ConradAlma)	0.497
1319	1Q (ButlerErew) - 1T (ConradAlma)	0.499
1320	1T (ConradAlma) - 2V (ForsterRoom)	0.500
1321	1J (MorrisDream) - 2G (OrczyElus)	0.503
1322	1E (ButlerErew) - 1T (ConradAlma)	0.508
1323	1K (StevensonCatri) - 1T (ConradAlma)	0.508
1324	1D (StevensonCatri) - 1T (ConradAlma)	0.511
1325	1T (ConradAlma) - 2W (StevensonCatri)	0.513
1326	1J (MorrisDream) - 1T (ConradAlma)	0.525

Appendix 3. Corpus Oxquarry 2

Author	Title	Length (tokens)	Vocabulary (types)
Braddon Mary	Fenton's Quest	172,952	10,154
Braddon Mary	The Golden Calf	160,172	12,490
Braddon Mary	London Pride	157,238	13,498
Braddon Mary	The Lovels of Arden	182,428	11,832
Braddon Mary	Phantom Fortune	178,994	12,886
Braddon Mary	Run to Earth	185,627	10,635
Butler Samuel	Erewhon	83,905	7,886
Butler Samuel	Erewhon Revisited	84,853	6,706
Butler Samuel	The Way of all Flesh	161,847	10,611
Chesterton G. K.	The Bell and the Cross	79,445	8,126
Chesterton G. K.	The Club of Queer Trades	44,017	5,663
Chesterton G. K.	The Innocence of Father Brown	78,902	8,169
Chesterton G. K.	Man Alive	57,553	7,492
Chesterton G. K.	The Man who was Thursday	58,066	6,494
Chesterton G. K.	Tremendous Trifles	55,969	6,821
Chesterton G. K.	The Wisdom of Father Brown	71,588	8,124
Conrad Joseph	Almayer's folly	62,575	6,483
Conrad Joseph	The Arrow of Gold	107,193	8,663
Conrad Joseph	Heart of Darkness	20,243	4,021
Conrad Joseph	Lord Jim	128,685	10,970
Conrad Joseph	Nostromo	170,103	12,805
Conrad Joseph	Typhoon	31,024	5,009
Conrad Joseph	Victory	117,319	9,408
Eliot George	Daniel Deronda	309,013	15,829
Eliot George	The Mill on the Floss	206,830	12,805
Eliot George	Middlemarch	316,156	16,369
Forster E. M.	Where Angels Fear to Tread	49,449	5,636
Forster E. M.	Howard's End	108,546	9,367
Forster E. M.	The Longest Journey	93,986	8,478
Forster E. M.	A Room with a View	66,456	6,920
Hardy Thomas	Beloved	62,839	7,191
Hardy Thomas	Jude	143,996	11,294
Hardy Thomas	Far from the madding crowd	137,879	11,958
Hardy Thomas	Wessex tales	80,694	8,231
Morris William	Dream	31,398	3,661
Morris William	News from nowhere	77,721	6,682
Orczy Baroness E.	The Elusive Pimpernel	83,314	8,114
Orczy Baroness E.	The Old Man in the Corner	69,591	6,365
Orczy Baroness E.	League of the Scarlet Pimpernel	75,208	7,151
Stevenson Robert L.	Ballantrae	87,156	7,694
Stevenson Robert L.	Catronia	101,012	7,940
Stevenson Robert L.	Doctor Jekyll	25,646	4,012
Stevenson Robert L.	Kidnapped	81,706	6,829
Stevenson Robert L.	Treasure Island	67,883	6,143
Tressel Robert	The Ragged Trousered Philanthropists	251,159	12,423
Wilde Oscar	The Picture of Dorian Gray	54,604	5,612
Wilde Oscar	Lord Arthur Savilles Crimes	39,311	5,351